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TRANSIENT EFFECTS IN M/G/1 QUEUES:

AN EMPIRICAL INVESTIGATION.

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PREPARED UNDER CONTRACT

NØØ014-76-C-0418

(NR-047-061)

NSF-ENG-75-14847 or the Office of Naval Research

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This research was supported in part by National Science Foundation Grant ENG 75-14847 Department of Operations Research, Stanford University issued as Technical Report No. 51

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CHAPTER 1

INTRODUCTION

1.1. Objective of the Dissertation

The objective of this dissertation is to provide tables for time-dependent expected server load in M/G/l queueing systems. The results are presented in a form that can be applied by practitioners involved in the design and control of operating systems. This presentation attempts to bridge the gap between the mathematical theory of stochastic processes and the numerical results useful to the management science practitioner.

Waiting lines and congestion are common problems encountered in almost everyone's daily life. A queue can develop in any service system whenever the immediate demand exceeds the system's capacity. One of the problems in designing or controlling such systems is achieving the proper balance between the level of service and the amount of waiting which might occur. Formal analysis may be appropriate if the service involves very expensive equipment, if the costs of waiting are extremely high, or if the decision involves many similar operating systems where the combined costs of service or waiting could be excessive.

The mathematical theory of queues can often provide some insight into the behavior of an actual or proposed operating system. Ideally,

This notation is the standard Kendall designation for queueing systems:

<u>Markovian</u> (or Poisson) arrivals/<u>General</u> service time distribution/<u>1</u>
(single) server.

the designer would prefer a mathematical model or algorithm that would determine the optimal system design, given the demand or arrival pattern, the costs of various service levels, and the costs associated with customer waiting or delay times. Unfortunately, there is no general optimization technique for queueing models. However, there are numerous predictive models that allow an analyst to determine some operating characteristics of a specific system. Prospective levels of service can be analyzed one at a time and an evaluative model can determine the total cost of service and waiting for each alternative.

The mathematical formulas for such operating characteristics that are often available to the practitioner apply only to queueing systems that are in statistical equilibrium. These steady state results are appropriate for a system where the server can, on the average, serve customers faster than they arrive and where sufficient time has passed to cancel the effects of the system's initial condition. If customers arrive at the system faster than the server can handle them, or if the behavior of the system must be analyzed for the start-up period, then the time-dependent (i.e., transient) operating characteristics are important. Lee [1966, p. 26], in an opinion probably shared by many practitioners, has stated "... the first working rule of queueing theory: time-dependent solutions to queueing-models are either unobtainable or unmanageable."

Transient solutions that are available in the queueing literature are usually expressed in terms of the Laplace transform of the operating characteristic. Bhat [1969, p. B284] reiterates the problem and mentions a possible solution:

"Plainly speaking, the results, given in terms of transforms, very often with more than one argument, fail to make sense to an applied researcher. Numerical inversion of transforms ... is an answer to this problem. But at this stage, the inversion methods are either not sophisticated enough to handle the more complex situations or do not appeal to the applied researcher."

The present research uses an accurate transform inversion technique to produce numerical results for transient expected system load. A practitioner can use these tables to analyze a wide range of M/G/1 systems by referring directly to Chapter 4. Chapter 2 discusses the transform relationships from which we start, and Chapter 3 discusses the inversion technique to be used.

1.2. Methodology

The operating characteristic studied in this disseration is time-dependent server load. The server load at epoch t, denoted W(t), is equal to the sum of the service times of customers waiting in the queue plus any remaining service time of a customer being served. The quantity W(t) is also called virtual waiting time, because it is the time that a hypothetical customer arriving at epoch t would have to wait before beginning service. Our objective is to tabulate the expected value of server laod, E[W(t)], for various epochs t and various system parameters (initial load, arrival rate, and service time distribution).

Our study of M/G/1 queueing systems can be embedded in the following general framework. Let $\{X(t), t \ge 0\}$ be a stochastic process

with stationary independent increments (a Lévy process) whose sample paths have no negative jumps. Let W(t) be this same process modified by a reflecting barrier at zero. If X(t) = S(t) - t, where $\{S(t), t \geq 0\}$ is a compound Poisson process with positive jumps, then W(t) is the server load process for an M/G/l queue. (The jumps of S(t) occur at customer arrival epochs and the jump sizes are service times.) We shall discuss two other choices for the X process which lead to W processes that provide bounds or approximations for the M/G/l server load process. In one instance, we take X(t) to be Brownian motion, and, in the other, we take X(t) = G(t) - t, where G(t) is a gamma process (a Lévy process whose increments are gamma distributed).

In Chapter 2 we present a Laplace transform result for E[W(t)] that covers both these two cases and the queueing (compound Poisson) case. In terms of its application to queueing processes, this result is valid for a system that begins operation either with or without an initial backlog of work and that has an average arrival rate less than, equal to, or greater than its average service rate.

There is no general closed-form expression for the exact inverse of the Laplace transform of mean server load, but numerical methods can be used to obtain an approximation of E[W(t)] at a specific epoch t. The numerical technique employed in this research computes E[W(t)] by taking a linear combination of the Laplace transform function evaluated at appropriate values of its argument. Theoretically, a more accurate approximation of E[W(t)] can be obtained by using more evaluations of

the Laplace transform expression, but, when using an electronic computer for the computations, its finite word length places a limit on the accuracy that can be achieved in the E[W(t)] approximation. Investigation of the technique using Laplace transforms with known inverses indicates that five or six significant digits for E[W(t)] can be obtained efficiently; this is more than enough to justify confidence in the three or four digit results shown in the final E[W(t)] tables.

The computational procedures used in this research could be applied to any M/G/1 queueing system whose service time distribution has a Laplace transform that can be evaluated numerically. In this research we achieve a balance between generality of results and ease of computation by concentrating on the well known class of M/G/1 queues whose service times have Erlang (or, equivalently, gamma) distributions. We denote these queueing systems $M/E_{k}/1$ and use the Erlang shape parameter k to describe the service times. (Parameter k is usually restricted to positive integer values when describing Erlang distributions, but here we allow k to assume any positive real value.) For example, the special case of k = 1 corresponds to the exponential service time distribution (an M/M/1 system). The present research also examines $M/E_{L}/1$ queues with k less than 1 and k greater than 1, corresponding to service time distributions with greater variance than the exponential and less variance, respectively. Most service time distributions encountered in actual practice can be adequately approximated using the very flexible Erlang family.

1.3. Overview of the Subsequent Chapters

In Chapter 2 we explain the sample path relationship between the server load process W(t) and the associated net input process X(t) in an M/G/l queueing system. In this discussion, the net input process can actually be any Lévy process which has no negative jumps. Breiman [1968] gives an introduction to the theory of such processes, and Blumenthal and Getoor [1968] provide a comprehensive treatment. We then present a result developed by Harrison [1977] for the Laplace transform of E[W(t)] when the net input is a general Lévy process. This theoretically oriented chapter finally examines a scaling procedure which facilitates comparisons among the various processes.

Chapter 3 explains the Laplace inversion technique developed by Gaver [1966] which gives approximation based on the expected value of an observational density function. The accuracy of Gaver's algorithm was improved by Stehfest [1970], and we test the technique using Laplace transform functions whose exact inverses are known. These numerical results are also compared with results from Veillon [1974] which were obtained using other Laplace inversion techniques. We then apply the technique to E[W(t)] in $M/E_k/1$ queueing systems, and we compare our results with Coleman [1975] who obtained exact E[W(t)] for the M/M/1 case by evaluating sums of Bessel functions. After some additional checks to ensure the accuracy of our approximations, we provide documentation of the computer programs which generate the tables of E[W(t)]. Gaver [1966, 1968] presented limited numerical results for transient

E[W(t)] in several M/G/l queues, thereby demonstrating the potential usefulness of his technique, and Coleman [1975] presented exact results only for the M/M/l case. The major contribution of the present research is the extensive set of tables in Appendix C, providing transient mean server load in queues with a wide range of traffic intensities, initial loads, and service time distributions.

Chapter 4 explains how the scaled results in the tables can be used by a practitioner to determine E[W(t)] in M/G/1 queues. Charts of the Erlang service time distributions are presented, and simple methods of interpolation in the tables are explained. Several sample problems demonstrate the entire procedure. It is intended that Chapter 4 can be used by a practitioner without recourse to Chapters 2 or 3.

The dissertation concludes in Chapter 5 with a brief summary, some qualitative observations, and suggestions for further research.

CHAPTER 2

THEORETICAL RESULTS FOR THE SERVER LOAD PROCESS

2.1. The Server Load Process in M/G/1 Queues

The M/G/l queueing system consists of a group of customers, a waiting room, and a service facility. We assume that customer arrivals are described by a stationary Poisson process $\{A(t); t \geq 0\}$ with mean arrival rate λ , where A(t) is the number of customers arriving during the time interval [0,t]. Thus, the probability of n arrivals in [0,t] is

$$P\{A(t) = n\} = \frac{e^{-\lambda t}(\lambda t)^n}{n!}$$
, $n = 0, 1, ...,$

and the times between consecutive arrivals are exponentially distributed with mean $1/\lambda$. We further assume that there are no bulk arrivals, no reneging, and no balking. The number of potential customers and the size of the waiting room are assumed to be unlimited, and the queue discipline is first-come-first-served.

Customers arrive at epochs $\{t_1, t_2, \ldots\}$, each with a demand for service $\{S_1, S_2, \ldots\}$. These individual customer service times are independent of the interarrival times and are independent, identically distributed non-negative random variables with finite mean E(S) and finite second moment $E(S^2)$. We assume that random variable S has a distribution function $F(\cdot)$ with corresponding Laplace-Stieltjes Transform (LST):

$$F*(s) = E(e^{-sS}) = \int_{0}^{\infty} e^{-sx} dF(x)$$
, $0 < s < \infty$.

The general approach used in this research requires only that this LST can be evaluated numerically. For the specific cases to be investigated, S has a continuous distribution with density function $f(\cdot)$, so the LST reduces to the ordinary Riemann integral

$$\int_{0}^{\infty} e^{-sx} f(x) dx .$$

In particular, we examine M/G/1 systems with gamma distributed service times, but we employ the terminology usually reserved for the Erlang family of distributions, where the two parameters are the mean E(S) and the shape parameter k. The Erlang density function is

(2.1.1)
$$f(x) = \frac{[k/E(S)]^k}{(k-1)!} x^{k-1} e^{-kx/E(S)}, \quad x \ge 0,$$

with $k \ge 1$ restricted to integer values. However, here we allow k to assume any non-negative real value, requiring that the factorial (k-1)! in the density function be replaced by the gamma function,

$$\Gamma(k) = \int_{0}^{\infty} x^{k-1} e^{-x} dx .$$

Using the Erlang terminology, the LST of our service time distribution is (see Drake [1967, p. 138] for a derivation)

(2.1.2)
$$F^*(s) = \left[\frac{k}{k + sE(s)}\right]^k$$
, $k \ge 0$,

The variance of an Erlang random variable is $[E(S)]^2/k$. Thus, for various Erlang service time distributions with the same mean, the shape parameter k is inversely proportional to the variability of those service times. For example, Erlang service times with 0 < k < 1 have even more variability than an exponential distribution (k = 1). On the other hand, for very large values of k the variance is very small, and we approach the case of a constant or deterministic service time (an M/D/1 system) as k tends to infinity. Probability density functions for $0 < k \le 1$ and $k \ge 1$ are shown graphically in Figure 4.1. Hillier and Lieberman [1974, p. 417] state that "empirical service-time distributions can usually be reasonably approximated by an Erlang distribution."

The most important descriptive parameter for a queueing system is its traffic intensity ρ , defined as the mean service time divided by the mean interarrival time. Thus, for M/G/1 systems, we have

 $\rho = \lambda E(S)$.

When ρ is less than one, i.e., when the average time between arrivals is greater than the average service time, we say that the system is "stable". In this case, ρ is the fraction of time that the server is busy, and it can be interpreted as the system's "utilization factor". Furthermore, for $\rho < 1$ the distribution of virtual waiting time

approaches an equilibrium distribution as t increases. Let W denote this steady-state waiting time. Its expected value is given by the Pollaczek-Khintchine formula,

(2.1.3)
$$E(W) = \frac{\lambda E(S^2)}{2(1-\rho)}$$
, $\rho < 1$.

Our numerical results for time-dependent virtual waiting time will allow us to observe how quickly this steady-state condition is approached. However, we will also examine "unstable" systems $(\rho \geq 1)$, where the queue length and waiting time tend to increase without bound, so that no steady-state conditions exist.

As mentioned in the introductory chapter, the virtual waiting time or server load, W(t), represents the work backlog at epoch t, i.e., the accumulated, unserviced demand. We define W(t) in terms of the underlying work input processes, which allows us to use existing Laplace transform results for reflected Lévy processes. Sample path relationships are shown graphically in Figures 2.1 and 2.2, illustrating one possible realization for these stochastic processes.

The input process S(t) represents the amount of customer work that arrives during the interval [0,t] and is defined by

$$S(t) = S_1 + \cdots + S_{A(t)}, \quad t \ge 0.$$

This compound Poisson process has mean $E[S(t)] = \rho t$ and variance $Var[S(t)] = \lambda E(S^2)t$.

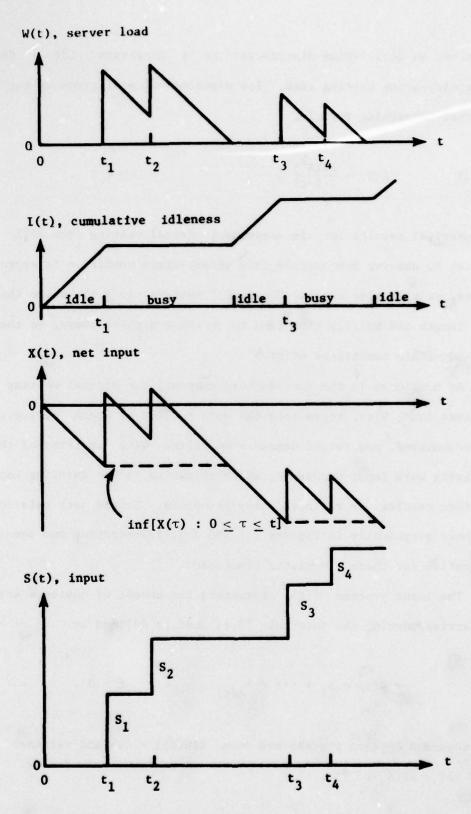
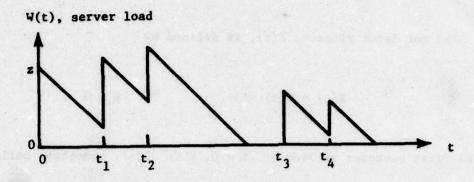
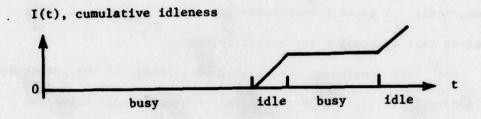


FIGURE 2.1. Sample Path Relationships when W(0) = 0.





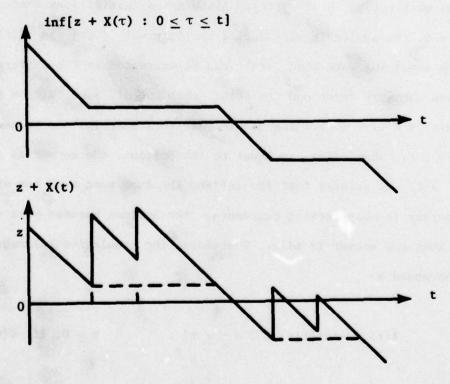


FIGURE 2.2. Sample Path Relationships when W(0) = z - 0.

The net input process, X(t), is defined by

X(t) = S(t) - t, $t \ge 0$.

If the first customer arrives at t = 0, then X(t), sometimes called pseudo-server load, is identical to the server load process as long as the server remains busy. The similarity ends when the server first becomes idle. A general expression for W(t) in terms of X(t) requires that we account for server idleness.

Let I(t) represent the cumulative idleness of the server during the interval [0,t]. Since the server works at a unit rate, the potential work output is t units during the same interval, and the actual work output is t-I(t). Assuming no initial work load, i.e., W(0) = 0, the available work during the interval [0,t] is S(t), so the remaining work load W(t) can be expressed as the difference between the work input and the actual work output, i.e., W(t) = S(t)-[t-I(t)], or W(t) = X(t) + I(t). The sample path relationships can be seen in Figure 2.1. When X(t) is equal to its infimum, the server is idle. When X(t) is greater than its infimum (because some work has arrived), the server is busy serving customers. The infimum becomes more negative only when the server is idle. Therefore, the cumulative idleness can be expressed as

 $I(t) = -\inf[X(\tau) : 0 \le \tau \le t]$, $t \ge 0$, if W(0) = 0.

Combining this with the net input, we have the server load representation (no initial workload),

$$W(t) = X(t) - \inf[X(\tau) : 0 \le \tau \le t]$$
, $t \ge 0$, if $W(0) = 0$.

The same reasoning applies to the more general case of initial server load, illustrated in Figure 2.2, where the underlying S(t) and X(t) sample paths would be identical to those shown in Figure 2.1. The initial server load z represents a specific amount of work available for service at epoch t=0. Then the available work during an interval [0,t] is z+S(t), and the remaining work load W(t) at any epoch t is z+S(t)-[t-I(t)], or W(t)=z+X(t)+I(t). The cumulative idleness function I(t) is slightly more complicated when W(0)>0. The server is initially busy in this case, so when z+X(t) is equal to its infimum, the server is idle only if that infimum is negative. After the initial busy period, the infimum of z+X(t) becomes more negative only when the server is idle, and the representation of the cumulative idleness function is similar to the case with no initial load. Thus, for this general case the server load representation is

$$W(t) = z + X(t) - \{\inf[z + X(\tau) : 0 \le \tau \le t]\}^{-},$$

$$t \ge 0, \text{ if } W(0) = z \ge 0.$$

Our main objective is to compute the expected value of W(t), where the expectation is taken with respect to a probability distribution over all possible sample paths. We adopt the notation $E_z[W(t)]$ and $E_z[I(t)]$ for expected server load and expected cumulative idleness, respectively, conditional on initial work load W(0) = z. From our discussion of the sample path relationships it follows that

$$E_{z}[W(t)] = E_{z}[z + S(t) - t + I(t)] ,$$

$$(2.1.3)$$

$$E_{z}[W(t)] = z + \rho t - t + E_{z}[I(t)] .$$

Thus, in M/G/1 queueing systems the mean server load can be separated into four additive terms: initial work load, new work input, potential work output, and cumulative idleness. Another useful observation is that $E_z[I(t)]$ must be zero in the interval from t=0 to t=z, i.e., it is impossible for the server to become idle before the initial work load has been serviced. In Chapter 3 this property is used to provide a check on the accuracy of our numerical results.

Thus far the sample path relationships and expectations have been discussed in the context of M/G/1 queueing systems. However, we also calculate $\operatorname{E}_{\mathbf{Z}}[W(t)]$ for processes where the same relationships apply, but where S(t) is not compound Poisson. These related processes provide bounds or approximations for M/G/1 mean server load.

2.2. The Laplace Transform Result for Expected Server Load

All of the stochastic processes for which numerical values are calculated here can be discussed simultaneously in the following unified framework. Let $X = \{X(t), t \ge 0\}$ be a process with stationary, independent increments (an infinitely divisible or Lévy process) and no negative jumps, and define

$$-\mathbb{E}[X(t)] = \mu t , \qquad \text{where } -\infty < \mu < \infty ,$$

$$(2.2.1)$$

$$\text{Var}[X(t)] = \sigma^2 t , \qquad \text{where } 0 < \sigma^2 < \infty ,$$

for $t \ge 0$. According to the standard results for Lévy processes, the Laplace transform of X(t) has the form

$$E[e^{-sX(t)}] = e^{-\Phi(s)t} \qquad \text{for } s > 0 \text{ and } t \ge 0,$$

and the exponent function $\Phi(\cdot)$ is convex with

(2.2.2)
$$\Phi(0) = 0$$
, $\Phi'(0) = \mu$, and $\Phi''(0) = \sigma^2$,

cf. Harrison [1977] and Takacs [1967]. It can be shown that for each s>0 there exists a unique $\omega(s)>0$ such that

(2.2.3)
$$\Phi[\omega(s)] = s$$
.

In the previous section we discussed the sample path relationships where process W is obtained from X by imposing a reflecting barrier at zero. Using the notation $E_z[W(t)] = E[W(t)|W(0) = z]$, let $P_W(z,s)$ denote the Laplace transform of $E_z[W(t)]$, that is,

$$P_{W}(z,s) = \int_{0}^{\infty} e^{-st} E_{z}[W(t)] dt$$
, $s > 0$.

Harrison [1977] developed a simple formula for the Laplace transform of $E_{\chi}[W(t)]$, where μ is unrestricted in sign:

(2.2.4)
$$P_{W}(z,s) = \frac{z}{s} - \frac{\mu}{s^{2}} + \frac{e^{-\omega(s)z}}{s\omega(s)}$$

In Chapter 3 we use this formula to obtain accurate approximations of $\mathbf{E}_{\mathbf{z}}[\mathbf{W}(\mathbf{t})]$ for specific epochs t, initial loads z, and various net input processes X. To achieve the desired accuracy the numerical inversion technique requires that $\mathbf{P}_{\mathbf{W}}(\mathbf{z},\mathbf{s})$ be computed for 34 values of s in order to obtain $\mathbf{E}_{\mathbf{z}}[\mathbf{W}(\mathbf{t})]$ at a single epoch, and each evaluation of $\mathbf{P}_{\mathbf{W}}(\mathbf{z},\mathbf{s})$ requires that the functional equation (2.2.3) be solved to obtain $\mathbf{w}(\mathbf{s})$.

The ease with which $\omega(s)$ can be calculated depends upon the form of the exponent function, and the exact form of $\Phi(\cdot)$ depends upon the process X. If we specify X(t) = S(t) - t, where S(t) is a compound Poisson process, then it can be shown that

$$\Phi(s) = s - \lambda[1 - F*(s)] ,$$

In the previous section we discussed the sample path relationships where process W is obtained from X by imposing a reflecting barrier at zero. Using the notation $E_z[W(t)] = E[W(t)|W(0) = z]$, let $P_W(z,s)$ denote the Laplace transform of $E_z[W(t)]$, that is,

$$P_{W}(z,s) = \int_{0}^{\infty} e^{-st} E_{z}[W(t)] dt$$
, $s > 0$.

Harrison [1977] developed a simple formula for the Laplace transform of $E_{z}[W(t)]$, where μ is unrestricted in sign:

(2.2.4)
$$P_{W}(z,s) = \frac{z}{s} - \frac{\mu}{s^{2}} + \frac{e^{-\omega(s)z}}{s\omega(s)}.$$

In Chapter 3 we use this formula to obtain accurate approximations of $\mathbf{E}_{\mathbf{Z}}[\mathbf{W}(\mathbf{t})]$ for specific epochs t, initial loads z, and various net input processes X. To achieve the desired accuracy the numerical inversion technique requires that $\mathbf{P}_{\mathbf{W}}(\mathbf{z},\mathbf{s})$ be computed for 34 values of s in order to obtain $\mathbf{E}_{\mathbf{Z}}[\mathbf{W}(\mathbf{t})]$ at a single epoch, and each evaluation of $\mathbf{P}_{\mathbf{W}}(\mathbf{z},\mathbf{s})$ requires that the functional equation (2.2.3) be solved to obtain $\mathbf{\omega}(\mathbf{s})$.

The ease with which $\omega(s)$ can be calculated depends upon the form of the exponent function, and the exact form of $\Phi(\cdot)$ depends upon the process X. If we specify X(t) = S(t) - t, where S(t) is a compound Poisson process, then it can be shown that

$$\Phi(s) = s - \lambda[1 - F*(s)] ,$$

cf. Prabhu [1965, p. 70] and Takacs [1967, p. 59]. Using the terminology of M/G/l queueing theory λ is the average arrival rate and F*(·) is the Laplace transform of the service time distribution. Since S(t) has mean $\lambda E(S)$ t and variance $\lambda E(S^2)$ t, it follows that the parameters of X defined by equations (2.2.1) are $\mu = 1-\rho$ and $\sigma^2 = \lambda E(S^2)$. If F*(·) can be evaluated numerically, then the properties (2.2.2) of $\Phi(\cdot)$ indicate that the functional equation $\Phi[\omega(s)] = s$ can be solved efficiently using an elementary one-dimensional search technique. In Chapter 3 we discuss our use of the Newton-Raphson method to obtain $\omega(s)$ for M/E_k/1 queueing systems.

For the special case of an M/M/l system the LST of the service time distribution is obtained by setting k=1 in equation (2.1.2), i.e., F*(s) = 1/[1+sE(S)]. For a specified value of s the functional equation $\Phi[\omega(s)] = s$ is a quadratic function of $\omega(s)$, and the positive solution is

(2.2.5)
$$\omega(s) = \frac{-\{1 - \lambda[s+E(S)]\} + \sqrt{\{1 - \lambda[s+E(S)]\}^2 + 4sE(S)}}{2E(S)}$$

The M/M/1 system is the only queue studied here that has an analytic solution to (2.2.3). In general the M/D/1 and M/E $_{\rm k}$ /1 cases will require a search to determine $\omega(s)$. The formulas for these cases are summarized in Table 2.3 below.

Two other choices for X(t) yield reflected processes W(t) which provide bounds or approximations for M/G/1 server load. The first case is the Wiener process, which has been used as a model for a variety of physical processes since its original development as a

Brownian Motion Process		1		$X_{W}(t) = \sigma \xi(t) - \mu t$ (ξ is standard	wiener process)	$\sigma^2 \mathbf{t}$	$\mu s + \frac{1}{2} \sigma^2 s^2$	n Quadratic Formula
98	M/D/1	u u	S_i constant $(k \rightarrow \infty)$	otali ori equipment alike ori	$(\lambda E(S) - 1)t$	λΕ(S) ² t	s-λ[1-e ^{-sE(S)}]	Newton-Raphson Search
M/G/1 Queueing Systems	M/M/1	$S(t) = S_1 + \cdots + S_{A(t)}$ where $A(t)$ is Poisson	S ₁ exponential (k = 1)	X(t) = S(t) - t	$(\lambda E(S) - 1)t$	λ2Ε(S) ² t	$s - \frac{sE(S)}{1 + sE(S)}$	Quadratic Formula
M/G/	$M/E_{\mathbf{k}}/1$	S(t) where	S ₁ Erlang with shape parameter k	100 S	$(\lambda E(S) - 1)t$	$\lambda \frac{k+1}{k} E(S)^2 t$	$s - \lambda \{1 - \left[\frac{k}{k + sE(S)}\right]^{k}\}$	Newton-Raphson Search
Gamma Input Process		$G(t)$ with parameters α, β	gamma distributed increments	$X_G(t) = G(t)-t$	$(\frac{a}{8}-1)t$	$(\frac{a}{\beta^2} - 1)t$	$s - a \ln(1 + \frac{s}{\beta})$	Newton-Raphson Search
		Input Process	Characteristics of the Input Process	Net Input Process X(t)	E[X(t)]	Var[X(t)]	Exponent Function $\Phi(s)$	Solution technique for $\Phi[\omega(s)] = s$

Summary of Characteristics for the Processes under Study. TABLE 2.3.

model for Brownian motion. Gaver [1968] proposed that this diffusion process could be used as an approximation for the net input process of an M/G/1 queue by equating the first and second moments of the two processes.

The Brownian motion process is denoted by

$$X_R(t) = \sigma \xi(t) - \mu t$$
,

where $\xi(t)$ is a Wiener rocess whose stationary, independent increments have a normal distribution with mean zero and unit variance. It follows that $X_B(t)$ has mean $-\mu t$ and variance $\sigma^2 t$. As previously noted, the net input process of an M/G/1 queue has mean $E[X(t)] = (\rho-1)t$ and variance $Var[X(t)] = \lambda E(S^2)t$. Thus, using $X_B(t)$ to approximate X(t) we would take $-\mu = \rho-1$ and $\sigma^2 = \lambda E(S^2)$. The exponent function for Brownian motion is $\Phi(s) = \mu s + \frac{1}{2} \sigma^2 s^2$, cf. Takacs [1967, p. 81]; thus the solution of $\Phi[\omega(s)] = s$ can be calculated using the quadratic formula.

The second non-queueing input process that we consider as an approximation is a gamma input process, denoted

$$X_{G}(t) = G(t) - t ,$$

where G(t) is a process whose stationary, independent increments have a gamma distribution. The gamma density function is

$$f(x) = \frac{\beta}{\Gamma(\alpha)} (\beta x)^{\alpha-1} e^{-\beta x}$$
, $\alpha > 0$, $\beta > 0$, $x \ge 0$,

so G(t) has mean $(\alpha/\beta)t$ and variance $(\alpha/\beta^2)t$. It follows directly that $X_G(t)$ has mean $(\alpha/\beta-1)t$ and variance $(\alpha/\beta^2)t$. Thus, using $X_G(t)$ to approximate the net input process of an M/G/1 queue we would choose α and β such that $\alpha/\beta-1=\rho-1$ and $\alpha/\beta^2=\lambda E(S^2)$. The exponent function for the gamma input process is $\Phi(s)=s-\alpha\log(1+s/\beta)$, cf. Takacs [1967, p. 66]; the solution of the functional equation (2.2.3) requires a search technique.

2.3. Scaling the Reflected Levy Processes

In this section we present a method for normalizing the processes of interest by using a simple linear tranformation for both the epochs and the server load. This technique allows us to reduce the tabulations required to describe actual operating systems and also facilitates comparisons among different processes by adopting a common, normalized time scale.

Consider an M/G/1 queueing system (as originally described in Section 2.1) with service times S_1, S_2, \ldots having distribution function $F(\cdot)$ with mean E(S) and second moment $E(S^2)$. Let $\{A(t); t \geq 0\}$ be the Poisson arrival process with mean arrival rate λ . The net input process is

$$X(t) = S_1 + \cdots + S_{A(t)} - t$$
,

and the server load process W(t) is obtained from [z + X(t)] by the reflection mapping described in Section 2.1, where z is the initial server load W(0).

We define a new scaled process in terms of the original queueing process

$$X^*(t) = aX(bt)$$
.

This scaled net input process has the form

(2.3.1)
$$X^*(t) = S_1^* + \cdots + S_{A^*(t)}^* - ct$$
, $t \ge 0$,

where c = ab, $A^*(t) = A(bt)$, and $S_i^* = aS_i$. Note that the random variables S_i^* have distribution function $G^*(x) = F(x/a)$, and that $A^*(t)$ is a Poisson process with mean arrival rate $\lambda^* = b\lambda$. Let μ^* and σ_{\star}^2 be defined by

$$E[X^{*}(t)] = -\mu^{*}t$$
,

(2.3.2)

$$Var[X^*(t)] = \sigma_{\bullet}^2 t .$$

We accomplish our normalization by choosing a and b so that $|\mu^{\bigstar}| = 1 \quad \text{and} \quad \sigma_{\bigstar}^2 = 1.$

By substituting aX(bt) for $X^*(t)$ on the left hand side of equations (2.3.2) and then using the parameters of process X as

defined by equations (2.2.1), it is easy to show that this particular scaling requires that

$$a = \frac{|\mu|}{\sigma^2}$$
 and $b = \frac{\sigma^2}{\mu^2}$.

If we define W^* as the reflection of $[z^* + X^*(t)]$, where the scaled initial server load is $z^* = az$, then it is easy to verify that $W^*(t) = aW(bt)$. That is, the reflection of the scaled net input process is identical to the scaled version of the original server load process, then it follows that

$$E_{z*}[W*(t*)] = \frac{|\mu|}{\sigma^2} E_{z}[W(\frac{\sigma^2}{\mu^2} t*)], \qquad z = \frac{\sigma^2}{|\mu|} z*.$$
In obtain scaled most

Thus, we can obtain scaled mean server load by evaluating an original queueing system and applying the transformation

$$\rho' = \lambda' E(S') = \lambda' \cdot \lambda E(S)/\lambda' = \lambda E(S) = \rho$$
,

defined by equations (2.2.1), it is easy to show that this particular scaling requires that

$$a = \frac{|\mu|}{\sigma^2}$$
 and $b = \frac{\sigma^2}{\mu^2}$.

If we define W* as the reflection of [z* + X*(t)], where the scaled initial server load is z* = az, then it is easy to verify that W*(t) = aW(bt). That is, the reflection of the scaled net input process is identical to the scaled version of the original server load process. If we let t* represent a (scaled) epoch for the scaled process, then it follows that

$$E_{z*}[W*(t*)] = \frac{|\mu|}{\sigma^2} E_{z}[W(\frac{\sigma^2}{\mu^2} t*)] , \qquad z = \frac{\sigma^2}{|\mu|} z* .$$

Thus, we can obtain scaled mean server load by evaluating an original queueing system and applying the transformations.

Before discussing the scaled process further, consider a second queueing system having Poisson arrival process A'(t) with mean rate $\lambda' \neq \lambda$ and service times S_1', S_2', \ldots with distribution function $F'(x) = F(\lambda' x/\lambda)$ so that $E(S') = \lambda E(S)/\lambda'$. The traffic intensity parameter for this second queueing system is the same as the original, that is,

$$\rho' = \lambda' E(S') = \lambda' \cdot \lambda E(S)/\lambda' = \lambda E(S) = \rho$$
,

but the variance parameter is different:

$$\sigma^{2} = \lambda' E(S^{2}) = \lambda' (\frac{\lambda}{\lambda'})^{2} E(S^{2}) = \frac{\lambda}{\lambda'} \lambda E(S^{2}) = \frac{\lambda}{\lambda'} \sigma^{2}$$
.

This second system and the original are related by

$$X'(t) = \frac{\lambda}{\lambda'} X(\frac{\lambda'}{\lambda} t)$$
,

or equivalently by

$$X'(t) = \frac{E(S')}{E(S)} X(\frac{E(S)}{E(S')} t) .$$

In the context of $M/E_k/1$ queueing systems, the second queueing system has the same ρ and same Erlang shape parameter k as the original, but the different values for σ^2 and ${\sigma'}^2$ require simple transformation of the epoch scale and waiting time scale.

Returning to the scaled process which was defined in terms of the original queueing process, we now express the original process in terms of the second process:

$$X^*(t) = \frac{|\mu|}{\sigma^2} X(\frac{\sigma^2}{\mu^2} t) = \frac{|\mu|}{\sigma^2} \frac{\lambda}{\lambda'} X'(\frac{\lambda'}{\lambda} \frac{\sigma^2}{\mu^2} t) = \frac{|\mu|}{\sigma'^2} X'(\frac{\sigma'^2}{\mu^2} t)$$

Once ρ (or μ) has been specified and a specific form of the service time distribution (i.e., a specific Erlang shape parameter k) has been determined, then we are free to choose the variance parameter (i.e.,

the time scales) for the queueing process from which X* will be obtained.

This research tabulates $E_{z*}[W*(t*)]$ for various values of ρ , k, and z*, and one can obtain mean server load for an actual queueing system by selecting the table with the closest $(\rho, k, z*)$ combination or by interpolating between two tabulated scaled processes, and then applying the transformation:

$$E_{z}[W(t)] = \frac{\sigma^{2}}{|\mu|} E_{z*}[W*(\frac{\mu^{2}}{\sigma^{2}}t)], \qquad z* = \frac{|\mu|}{\sigma^{2}} z.$$

Since $\mu=1-\rho$, this particular normalization cannot be used for systems with $\rho=1$, i.e., $\mu=0$. In this case we tabulate mean server load for Erlang queueing systems with λ and E(S) chosen so that the variance parameter $\sigma^2=\lambda E(S^2)=1$. Let superscript zero indicate parameters for the tabulated system. Then $\sigma^2=1$ requires that

$$\lambda^0 = \frac{k+1}{k}$$
 and $E(S^0) = \frac{k}{k+1}$.

Mean server load for an actual operating system with shape parameter $\,k\,$ and mean arrival rate $\,\lambda\,$ can be obtained from the tabulated values as follows:

$${\rm E}_{\rm z}[{\rm W}({\rm t})] = \frac{\lambda^0}{\lambda} \; {\rm E}_{\rm z0}[{\rm W}^0(\frac{\lambda}{\lambda^0} \; {\rm t})] = \frac{k+1}{k\lambda} \; {\rm E}_{\rm z0}[{\rm W}^0(\frac{k\lambda}{k+1} \; {\rm t})] \; , \quad {\rm z}^0 = \frac{k\lambda}{k+1} \; {\rm z} \quad .$$

The tabulated values for the ρ = 1 case correspond to an actual queueing system, whereas the tabulated values for the ρ ≠ 1 cases represent a stochastic process with "potential work output rate" of $c = ab = 1/|\mu|$, interpreted from equation (2.3.1). For the ρ < 1 cases, mean server load in queueing systems approaches the Pollaczek-Khintchine values, equation (2.1.3). Thus, the tabulated values for all scaled processes approach

$$\frac{|\mu|}{\sigma^2} \cdot \frac{\sigma^2}{2|\mu|} = \frac{1}{2} \quad ,$$

thereby allowing comparisons concerning approach to equilibrium.

CHAPTER 3

NUMERICAL INVERSION OF THE LAPLACE TRANSFORM

3.1. Approximation by an Expected Value

In the previous chapter we cited an expression for the Laplace transform for the three processes of interest. For the two special cases of Brownian motion and the M/M/l queue, analytic methods can be applied and exact numerical solutions can be obtained. However, in general it is necessary to use a method for numerical inversion of the Laplace transform in order to evaluate the process of interest. In this chapter we describe such a method and its implementation.

The first two sections of the chapter describe the method for approximate transform inversion using an expected value. In the third section we present numerical results for test cases where this method is applied to Laplace transforms whose exact inverses are known, and we describe the computer routines that apply the inversion method to prepare tables of E[W(t)].

The method used in this research for numerical inversion of the Laplace transform was originally developed by Gaver [1966]. Although there are other techniques which could have been implemented, this particular method was chosen because it had been applied successfully to the Laplace transform expression for E[W(t)] in the M/M/l and M/G/l queues by Gaver [1966, 1968].

The problem of inverting the Laplace transform can be summarized as follows. We have a function of interest, P(t), for which no closed-form analytic expression is known, so that it cannot be evaluated

numerically. However, we do know the expression for its Laplace transform,

$$p(s) = \int_{0}^{\infty} e^{-st} P(t) dt .$$

Specifically, in this research the function or process of interest is E[W(t)] and the Laplace transform expression was cited in Chapter 2. In general, we wish to tabulate P(t) for various values of t.

The method presented here computes an approximate value of the function at a specified point t'. Theoretically this approximation can be computed as precisely as desired. Consider observing the function of interest at a random time T that has density function f(t) such that the values of T are concentrated near t', as shown in Figure 3.1. Then $\bar{P}(t')$, an approximation of P(t'), can be expressed as

(3.1.1)
$$\overline{P}(t') = \int_{0}^{\infty} P(t) f(t) dt.$$

Now the problem is one of finding an observational density function f(t) so that the right hand side of (3.1.1) can be expressed in terms of p(s).

Gaver [1966] examined such a family of density functions that are well-suited to numerical computation. Let $f_n(t; a)$ be the density function for random variable T, with parameters n and a, as follows:

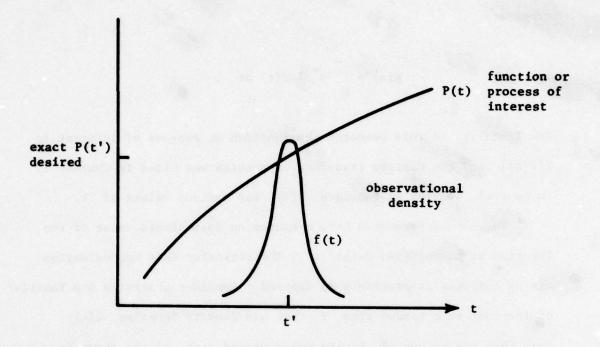


FIGURE 3.1. Observational Density Function

3.1.2)
$$f_n(t; a) = a \frac{(2n)!}{n! (n-1)!} (1 - e^{-at})^n e^{-nat}, a > 0, n = 1,2,...$$

Gaver [1966] showed that $f_n(t; a)$ has the following properties:

modal value
$$\approx \frac{1}{a} \ln 2$$

$$Var(T) \approx \frac{1}{a^2} \frac{2n+2}{(2n-1)(4n+1)}$$
.

For increasing values of n, $f_n(t; a)$ becomes more sharply peaked at $t = 1/a \ln 2$. Using $f_n(t; a)$ as the observational density function

in (3.1.1), we let the approximation of P(t') be

(3.1.3)
$$\bar{P}_n = \int_0^\infty P(t) f_n(t; a) dt$$
,

where parameter a is chosen such that a = 1/t' ln 2, thereby concentrating the values of random variable T near t' as desired.

By substituting the observational density (3.1.2) in the approximating expression (3.1.3) and then expanding $(1 - e^{-at})^n$ by the binomial theorem, we obtain

(3.1.4)
$$\bar{P}_n = a \frac{(2n)!}{n! (n-1)!} \sum_{i=0}^n {n \choose i} (-1)^i p[(n+i)a]$$
.

Thus, \overline{P}_n , an approximation of P(t'), is based on a linear combination of the Laplace transform p(s) evaluated at n+1 different values of s. Theoretically, the sequence $\{\overline{P}_n; n=1,2,3,\ldots\}$ converges to P(1/a ln 2), i.e., to P(t'). For any finite n, we can calculate \overline{P}_n using (3.1.4), and because $Var(T) \to 0$ as $n \to \infty$ we can obtain as good an approximation as desired by using a large enough value of n. Practically, p(s), the Laplace transform of P(t), must be evaluated at n+1 values of s in order to evaluate \overline{P}_n ; in some applications much time may be required to make each evaluation of the Laplace transform expression and so the cost of obtaining the desired precision in \overline{P}_n may be prohibitive. Also, for large n the factorial terms in (3.1.4) are integers with too many digits to be precisely evaluated on computers with finite word length, and operations involving these large numbers

can result in considerable rounding errors. From the standpoint of computational efficiency it is important to improve the accuracy of the approximate inverse without requiring evaluation of \bar{P}_n for extremely large values of n.

3.2. Improving the Accuracy of the Approximation

Gaver [1966] showed that each \overline{P}_n can be represented as an asymptotic expansion, i.e.,

(3.2.1)
$$\bar{P}_n \approx P(t') + \frac{\alpha_1}{n} + \frac{\alpha_2}{n^2} + \frac{\alpha_3}{n^3} + \cdots$$

where the error components α_i depend only on t' and not on n. An improved approximation of P(t') can be obtained by taking a linear combination of a set of the \overline{P}_n approximations, where the values of n are integer powers of 2. This method, extrapolation to the limit, ensures that some of the error terms in (3.2.1) are cancelled out, cf. Gaver [1966] and Henrici [1964].

Stehfest [1970] demonstrated an improved calculation method that uses a linear combination of an even number of the \overline{P}_n approximations; this method cancels even more of the error terms than Gaver's method (extrapolation to the limit). Stehfest combined the coefficients for the linear combination of \overline{P}_n with the coefficients of p(s) in (3.1.4) so that F_n , the approximation of P(t'), could be expressed directly as a linear combination of p(s):

(3.2.2)
$$Fa = \frac{\ln 2}{t!} \sum_{i=1}^{N} V_i p(\frac{\ln 2}{t!} i) .$$

Here N must be even and the combined coefficients V, are

(3.2.3)
$$V_{i} = (-1)^{(N/2)+i} \sum_{k=\left[\frac{i+1}{2}\right]}^{Min(i,\frac{N}{2})} \frac{k^{N/2} (2k)!}{(\frac{N}{2}-k)! \ k! \ (k-1)! \ (i-k)! \ (2k-1)!}.$$

For a theoretical comparison of the original \bar{P}_n , extrapolation to the limit, and \bar{P}_n consider the three methods where the number of values of p(s) is 32 in each case. Thirty-two evaluations of the p(s) could be used to compute \bar{P}_{31} , and from (3.2.1) the first error term would be $a_{31}/31$. On the other hand the thirty-two values of p(s) could be used to compute \bar{P}_1 , \bar{P}_2 , \bar{P}_4 , \bar{P}_8 , and \bar{P}_{16} , and it could be shown that extrapolation to the limit would yield an approximation of p(t') where the first four error terms of equation (3.2.1) would cancel completely. The resulting approximation would be $p(t') + a_5/2^{10} + \cdots$. Finally using Stehfest's method with $p(t') + a_1/2^{10} + \cdots$. Finally using Stehfest's method with $p(t') + a_1/2^{10} + \cdots$. This theoretical superiority of Stehfest's method was verified numerically in preliminary investigations using functions with known inverses.

3.3. Computer Implementation of the Technique

In this research, the algorithm LINV developed by Stehfest [1970] was translated from ALGOL to FORTRAN IV. Some changes were made in evaluating the V_i of equation (3.2.3) to take advantage of cancelling the factorial terms and thereby avoid rounding errors. Some preliminary numerical investigations used BASIC and single and double precision FORTRAN IV. The computer programs listed in Appendix B specify double precision for all non-integer variables, and the "AUTODBL" option available with the IBM FORTRAN H-Extended Compiler was used in all of the research reported here. With this extended precision there are approximately 36 significant decimal digits for the non-integer variables.

Stehfest [1970] published a table of results applying LINV to six transforms whose inverses are known, using 8 digit arithmetic and n = 10. Figure 3.2 shows the exact values of the six functions, the approximations from this research using N = 34 with 36 digit arithmetic, and the original Stehfest approximations using N = 10. We observe that there are at least six correct significant digits in our approximations (using N = 34) for each of the test functions.

Since the LINV approximation (Fa) is based upon the original \bar{P}_n , it should also be more accurate when more values of the Laplace transform p(s) are used, i.e., if N is very large. However, for large values of N the rounding errors in evaluating V_i can affect the accuracy of the results. This problem was investigated by applying LINV to the six test functions of Figure 3.2; approximations were obtained using

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FIGURE 3.2 INVERSION OF TEST FUNCTIONS

		MATE F(T) US ING		MATE F.	T) USING ETHOD
-	EXACT F(T)	40 = Z	X 110	EXACT F(T)	4 M II Z	011
	F(T)	= 1/SORT(PI*T)		F(T)	= -C-LN(T)	
1	444000		2073			00000
•	#0041 4000 #000	•364169363	• 5055	0.5//2156/1	0.5//215004	9118
•	· 398942259	• 35 894ZZ 80	. 3991	1.270362852	1.270362845	1.2708
•	.325734991	•325735007	• 3265	1.675827960	1.675827953	1.6754
	.282094777	. 282094791	.2827	1.963510032	1.963510026	1.9639
	•252313239	•252313252	. 2517	2.186653584	2.186653577	2.1872
	.230329421	•230329433	. 2298	2.368975140	2.368975134	2.3687
	•213243607	•21 324361 8	. 2132	2.523125820	2.523125814	2.5227
	.19947112	.199471140	. 1995	. 656657213	•656657206	.6574
	.188043184	.188063194	. 1881	2.774440248	2.774440242	2, 7739
10.0	8412402	12	0.17796	2.87980076	-2.8798007579	-2.88091
	-	= (T*T*T)/6		F(T)	= EXP(-T)	
	166666666	6666666	1656	367879441	367879441	3679
	33333333	33333333	3254	135335283	.135335283	1355
	-5000000000	50000000	4735	049787068	.049787068	-0504
	0.6666666666	0.666666666	0.6034	018315638	.018315638	.0184
	0.83333333	0.83333333	0. 7084	006737947	.006737947	.0064
	00000000009	0000000009	5. 7883	002478752	. 002478752	.0019
	7.166666666	7.166666666	6.8253	000911882	•000911882	• 0003
•	85, 333333333	85.23333333	84.8273	000335462	.000335462	000000
000	0	121 -500000000000	120.78473	0.0001234098	0.0001234097	-0000047
		000000000000000000000000000000000000000		***************************************	***************************************	
	F(T	N (SORT (2 *T		F(T)	*T*T/	9
	997765	9677659	. 987	• 66666666	9999999	•665
•	.909297426	.909297474	.9100	0.3333333	0.33333333	0.3253
	638157635	•638157668	• 6382	000000000	00000000	.0257
•	0.308071742	0.308071758	0.3096	• 33333333	33333333	• 3953
•	10000000000000000000000000000000000000	256586020-0	1170.0	00000000		* 1884
•	0.510947105	# 114601000000000000000000000000000000000	0. 51 VC	3.6666666666666666666666666666666666666	000000000	1022 -
	756802405	264693920 766802534) (12, 313,33	12-111111	11 8205
	457249168-0	0-891682300	910	00000000	6 000000000	5.2830
	0.97127759	0.971277849	9686 0	45.66666666	45.66666666	44.8851

N equal to 30, 32, 34, and 36. In each case the correct number of decimal places was recorded, rounding off both the exact value and the approximation when making the comparison. Figure 3.3 shows the average number of correct decimal places over the ten values of t evaluated in each case. We observe that the approximations using N = 34 are at least as good as those using N = 32 or N = 36 in all cases. Thus, in the remainder of this research we will use N = 34 in the LINV algorithm.

Figure 3.4 compares the LINV algorithm with other inversion techniques. The top section of this table reproduces some results by Dubner and by Veillon (1974)²; Dubner's approximations are based on 500 values of p(s), while Veillon's method uses 64 values. We observe that in all cases the LINV algorithm using 34 values of p(s) is at least as accurate as their techniques. The bottom section of Figure 3.4 compares approximations of a second function; we observe that once again the LINV results are as accurate as the results obtained by other techniques.

Figures 3.2, 3.3 and 3.4 indicate that LINV can provide very accurate approximations for the test functions, but it is also important to investigate LINV's accuracy using the function of interest in this

In preliminary investigations using FORTRAN IV double precision (approximately 17 digit arithmetic) without the AUTODBL option, the most accurate approximations were obtained using N = 18.

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FIGURE 3.3. Average Number of Correct Rounded-off Decimal Places in LINV Approximations of Six Functions for Determination of Optimal Value of N

N, the number of values of the Laplace transform p(s) used by LINV to estimate the function

Function	N = 30	N = 32	N = 34	N = 36
$\frac{1}{\sqrt{\pi t}}$	6.9	6.9	6.9	6.9
$\sqrt{\pi t}$ $\frac{t^3}{6}$	9.7	10.5	12.7	11.9
sin √2t	6.7	6.7	6.7	6.7
- C - 1n t	7.3	7.3	7.5	7.4
e ^{-t}	9.7	10.1	10.5	10.0
$1 - 3t + \frac{3t^2}{2} + \frac{t^3}{6}$	9.7	10.5	11.9	10.6

FIGURE 3.4 COMPARISONS WITH OTHER TECHNIQUES

F(T) = 1/	1/50FT(T*P1)	PROXIMAT	F(T) US ING	UMERICAL	I NVE RS 10
		STEHFEST	THOD		
H	XACT F(T)	AH NA	12	DUBNER*	VEILLON*
:		i	-	-	
•	.5641855	. 5641895		0. 73172	. 5641
	.3589422	. 3989422	.399	.4003	. 3989
	.3257349	. 3257350	.326	.2634	.3257
•		N	282	828	0
	. 25 231 32	.2523132	.251	.2936	.2523
	.2303294	. 230 3294	. 229	.2290	. 2303
	•2132436	.2132436	.213	.1806	.2132
	.1994711	. 1394711	.199	. 2011	.1994
	.1860631	.1880631	.188	2160	.1880
	•17E4124	. 1784124	•177	.1765	.178
F(1) = EX	F(-1/2)		IMATE F(T) US	ING NUMER!	ICAL INVERSION
-	EXACT	ST	EMFEST N=34	BELLMAN*	VE ILL ON
.14018	26174	74277 0.1	17402742	05	0.126
. 50112	•286343	86283 0.2	343498628	.28819	0.28632
. 54343	•439675	82721 0.4	675128271	• 43908	0.43967
.08508	.581268	87062 0.5	268 79 87 06	.58130	0.58126
•69314	• 70 71 66	37334 0.7	106843733	.70731	0.70710
.41229	.813711	45453 0.8	711814545	. 81340	0.81371
• 21482	•89 E1 56	24175 0.8	156912417	. 89848	0.89815
14000000	0.95813125	12155 0.9	5813125150	• 95784 00000	0.95813
0010	000766	600 67710	77 1901800	. 33220	10266.0

CATA FOR METHODS MARKED BY AN ASTERISK WERE TAKEN FROM A.C.M ALGORITHM 4.86. "NUMERICAL INVERSION OF LAPLACE TRANSFORM", BY FRANCOISE VEILLON, COMMUNICATIONS OF THE A.C.M., VOLUME 17. NUMBER 10. *NOTE:

research, E[W(t)]. Our earlier investigations using LINV (not included here) verified the results for M/M/1 queues reported by Gaver [1968], but it is more appropriate to compare our LINV results with mean server load obtained by some method other than Laplace transform inversion. Fortunately, Coleman [1975] evaluated E[W(t)] in the M/M/1 queue using a sum of Bessel functions. In Figure 3.5 the first three columns show his results at five epochs and the LINV results using equations (2.2.4) and (2.2.5) for the Laplace transform expression. The five significiant digits available in the Coleman results are matched exactly when the LINV results are rounded off.

As mentioned in Section 2.2, except for the Wiener process and the M/M/l queue, the value of $\omega(s)$ must be obtained using a search technique. Our final results use the quadratic formula to determine $\omega(s)$ for the Wiener process, but in all other cases (M/D/l, M/E $_k$ /l including M/M/l, and the gamma input process) we use the standard Newton-Raphson method. This search technique was chosen because expressions for the derivatives of the functions were available and subsequent computations demonstrated that the search usually converged in only four or five iterations.

The exponent function $\Phi(\cdot)$ is designated FMD1, FMEK, or FGAM in the FORTRAN subroutines, and functions FOMD1, FOMEK, and FOGAM are defined as

FO ···
$$[\omega(s)] = F ··· [\omega(s)] - s$$

Thus, in order to determine $\omega(s)>0$ such that $\Phi[\omega(s)]=s$, we search for the value of ω such that FO \cdots equals zero. Subroutine ZERO performs the search, computing the relative accuracy (change in ω from the last iteration). When the absolute value of the relative accuracy is less than a specified tolerance, the search stops, and the current value of ω is used to compute the Laplace transform using equation (2.2.4).

The choice of tolerance value (FORTRAN variable TOLRNC) affects the accuracy of ω determined from the search, and the accuracy of ω affects p(s) and the estimate of E[W(t)]. Figure 3.5 shows both the approximations of E[W(t)] based upon the exact quadratic solution of the functional equation (2.2.4) and the approximations obtained using the Newton-Raphson search with various specified tolerances. We observe that the two results agree for nine or ten significant digits in all cases and that the number of search iterations does not increase greatly as the tolerance is decreased. For each of the ten epochs, ω (s) is evaluated thirty-four times (N = 34); thus, the total number of iterations shown in Figure 3.5 applies to 340 evaluations of ω (s). Since the average number of iterations in each evaluation only increases from 4.2 to 5.3 as the tolerance decreases from 10^{-10} to 10^{-24} , we have set the search tolerance at 10^{-20} .

In Chapter 2 we discussed the stochastic process W(t) in an M/G/1 queue and its expected value function. Equation (2.1.3) is restated here in a slightly different form along with the Laplace transform of $E_{\sigma}[W(t)]$, equation (2.2.4):

Thus, in order to determine $\omega(s)>0$ such that $\Phi[\omega(s)]=s$, we search for the value of ω such that F0 ··· equals zero. Subroutine ZERO performs the search, computing the relative accuracy (change in ω from the last iteration). When the absolute value of the relative accuracy is less than a specified tolerance, the search stops, and the current value of ω is used to compute the Laplace transform using equation (2.2.4).

The choice of tolerance value (FORTRAN variable TOLRNC) affects the accuracy of ω determined from the search, and the accuracy of ω affects p(s) and the estimate of E[W(t)]. Figure 3.5 shows both the approximations of E[W(t)] based upon the exact quadratic solution of the functional equation (2.2.4) and the approximations obtained using the Newton-Raphson search with various specified tolerances. We observe that the two results agree for nine or ten significant digits in all cases and that the number of search iterations does not increase greatly as the tolerance is decreased. For each of the ten epochs, $\omega(s)$ is evaluated thirty-four times (N = 34); thus, the total number of iterations shown in Figure 3.5 applies to 340 evaluations of $\omega(s)$. Since the average number of iterations in each evaluation only increases from 4.2 to 5.3 as the tolerance decreases from 10^{-10} to 10^{-24} , we have set the search tolerance at 10^{-20} .

In Chapter 2 we discussed the stochastic process W(t) in an M/G/1 queue and its expected value function. Equation (2.1.3) is restated here in a slightly different form along with the Laplace transform of $E_2[W(t)]$, equation (2.2.4):

FIGURE 3.5 EFFECT OF NEWTON-RAPHSON SEARCH TOLERANCE

LAPLACE AND THE TOLERANCES. THIS TABLE COMPARES CCLEMAN'S BESSEL FUNCTION RESULTS AT FIVE EPOCHS WITH INVERSION RESULTS FOR BOTH THE EXACT QUADRATIC SOLUTION OF THE FUNCTIONAL APPROXIMATE NEWTON-RAPISON SEARCH SULUTION OF THE FUNCTIONAL WITH VARIOUS

AVM/1 MEAN SERVER LOAC. LAMEDA = 1.00. E(S) = 0.95. Z = 0.

7/4/1	MEAN SERVER	LOAC. LAMED	= 1.00 E(S) =	•0 = Z • S6•		
			SION OF LAPLAC	TRANSFORM. ST	FEST'S METHOD	4 M I N
FEDCH	N N	ט ע	TOLERANCE FOR	NEWTON-RAPHSON	SEARCH SOLUTION	OF FUNCTIONAL
	COLEMAN	FUNCTION	10	10**-12	10**-1	10**-16
20.	1	180512301	9180512301	01 80 51 2 30	10180512301	105512301
40.		46	5.46511190320	5.46511190320	5-46511190320	5-46511190320
•09	58	• 55e181 00 63	. 5581810063	. 5581810063	.5581810063	. 5581810063
80.	0	• 41 9C89 581 5	• 41 90899815	•4190899815	•4190899815	.4190899815
0	33	• 133€43204€	.1335432045	.1335432045	.1335432045	.1335432045
0		0.5687887379	0.5687887379	0.5687887379	0.5687887379	0.5687887379
0		2.0815626053	2. 0819626053	2.0819626053	2.0819626053	2.0819626053
0		3-1502768763	.1502768763	.15 (2768763	.1502768763	.1502768763
0		5.4923083051	5.4 323083091	5.49 230 83091	5.4923083091	5.4923083092
1 200		• 56 15599606	6.561959609	6. 5619599605	6 561959609	6.5619599606
6	MBER CF	NE STON-RAFHSCN				
TE	10 E	LUATE 10 EPOC	HS: 1423	1460	1518	1602
	9	INVER	IN OF LAPLAC	SFORM. ST	HFEST'S M	N=34
2	BESSEL	CUACRA	OLERANCE FO	N-RAPHSON	SEARCH SOLUT	OF FUNCTION
-	COLEMAN	FUNCTIONAL	10**-	10**-20	**-22	10**-24
	0	1026132010	1010101010	102010	2010010	1056190016
	246	100710016	10 63 1606 16	1057150016	1067160016	1057150016
	0 4	25 06 111 20 4 B	5591 81 00 62 5581 81 00 62	2506111699	25061116040	25061115040
	400	A100806815	410080815	41 90 80 91 E	410080015	419089915
C	8-1335	1335432045	1335432045	1335432045	1335432045	1335432045
0		0.5667887379	10.55878873791	10.56878873791	0.5687887379	10.56878873791
00		2.0815626053	2.0819626053	2,0819626053	2,0819626053	2.0819626053
0		3.1502768763	3-1502768763	3.1502768763	3-1502768763	3.1502768763
0		5.49 2 308 30 51	. 4 92 30 8 30 91	. 4923083091	. 4923083091	•4923083092
1200	1	16.56195996067	6.5619599606	6.5619599606	6.5619599606	6.5619599606
410	NUMBER OF	NUMBER OF STREET				
EK	0	UATE 10 EPCC	1710	1756	1773	1671

$$E_{z}[W(t)] = z - (1-p)t + E_{z}[I(t)]$$
,

$$P_{W}(z,s) = \frac{z}{s} - \frac{\mu}{s^{2}} + \frac{e^{-\omega(s)z}}{s(s)}$$
,

The first two terms of $P_{W}(z,s)$ can be inverted "by inspection," i.e., by referring to any table of function-transform pairs. Only the last term requires numerical inversion using LINV. Stehfest [1970, p. 48] cautioned the prospective user of LINV as follows: "One ought to be sure a priori that the unknown function F(t) has not any discontinuities, salient points, sharp points, sharp peaks, or rapid oscillations." Recall that our unknown function, $E_{\sigma}[I(t)]$ in M/G/1 queues, is a nondecreasing function and that it must be zero between t = 0 and t = z. Although it is not a discontinuous function, its slope does change abruptly at t = z, and we might expect some inaccuracies near that point. The top section of Figure 3.6 shows scaled approximations of $E_{z}[I(t)]$ evaluated at t = z for M/M/1 queues with traffic intensities between .5 and 2 and scaled initial server loads between .2 and 4. If the inversion technique is accurate, then all entries in the table should be zero. We observe that there are some inaccuracies, particularly for queues with low traffic intensities. For example, the worst case is ρ = .5 with z' = 1.0, where scaled E_z[I(t)] is .007 instead of zero.

Since the inaccuracies may be related to the discontinuity of the derivative of $E_z[I(t)]$ at t=z, our approach does not invert the third term of $P_W(z,s)$ directly. Instead, we construct a new function,

FIGURE 3.6 CORRECTIONS FOR DISCONTINUOUS FIRST DERIVATIVE

APPRCXIMATIONS USING STEHFEST'S METHOD. N=34

M/M/1 SCALEC MEAN CUMULATIVE IDLENESS AT T"=ABS(MU) #2" (I.E. AT T=2)

WITHOUT CORRECTION TERM

SCALED INITIAL SERVER LOAD

FHO	0.2			0.8	1.0	2.0	3.0	•••
!							*******	
.5	1500	0.005112	0-006274	00684	0000	0051	.00271	0.001107
9.	0026	0.003710	0-003879	00 360	0031	6000	000010	000003
1.	0.001984	0.002061	0.001606	0-001108	0.000000	400000	000021	000000
8.	00100	0.000587	0.000235	00000	0000	0000	00000	000000
6.0	0000	000 000 00	00 0005	00000	0000	- 000000	0000	000000
-	0000	100000	000000	00000	0000	0000	00000	00000000
2	0000	0.000022	000000	00000	0000	0000	0000	00000000
.3	4000	0.000083	0.000011	00000	0000	0000	00000	00000000
*	0000	0.000151	0.000000	00000	0000	0000	00000	000000
.5	9000	0.000209	0.000051	00000	0000	0000	00000	-000000
0.	0.000771	0.000313	0.000095	0000	0000	0000	00000	00000000

M/M/1 SCALED MEAN CUMULATIVE IDLENESS AT T"=ABS(MU) +2" (I.E. AT T=2)

WITH CORRECTION TERM

			S		SER	AD		
RHO	0.2	0.0	9•0	0.8	1.0	2.0	3.0	0.4
-								
0.5	0.00002		000001	000000	000010	000033	000129	0 0 0 2 8 9
9.0	0.000002		-000000	000000	000010	000016	000157	000150
0.7	0.00000		+000000-		000020	000000	000024	0.00000
0.8	000001		000011	000019	000021	- 000001	00000000	- 000000
6.0	000001		000003	- 000000	00000000	- 000000	-000000	000000-
1.1	000002		-000000	00000000	0000000	00000000	- 000000	00000000
1.2	00000000		- 000005	- 000001	-0000000	- 000000	00000000	0000000
1.3	0.000000		000003	000005	1000000-	00000000	- 000000	00000000
1.4	0.00000		-,000002	000005	000001	00000000	00000000	- 000000
1.5	0.000000		100000	- 00 0000	- 000001	000000-	00000000	-000000
2.0	0.00000	0 • 000 005	0.00000	-• 000000	000001	- 000000	000000	00000000

$$H(t) = E_{z}[I(t)] + g(t)$$
,

where g(t) and its Laplace transform are known, and where g(t) is chosen so that both H(t) and its derivative are continuous. Then LINV is used to invert the Laplace transform of H(t), and approximate $E_z[I(t)]$ is obtained by subtracting g(t) from approximate H(t).

It can be shown that the slope of $E_z[I(t)]$ in M/G/1 queues is equal to the probability that the server load is zero, denoted $P\{W(t)=0\}$, an any epoch t. Therefore, the slope of $E_z[I(t)]$ at t=z is $P\{W(z)=0\}$, which is the probability that no arrivals will occur between t=0 and t=z, or $e^{-\lambda z}$. For H(t) to have a continuous first derivative at t=z, that derivative must be zero. Since the slope of $E_z[I(t)]$ is $e^{-\lambda z}$ at that point, then the slope of $E_z[I(t)]$ is $e^{-\lambda z}$ at that point, then the slope of $E_z[I(t)]$ is $E^{-\lambda z}$ at that point, then the slope of $E_z[I(t)]$ is $E^{-\lambda z}$ at that point, then the slope of $E_z[I(t)]$ is $E^{-\lambda z}$ at that point, then the slope of $E_z[I(t)]$ is $E^{-\lambda z}$ at that point, then the slope of $E_z[I(t)]$ is $E^{-\lambda z}$. Therefore, we define $E_z[I(t)]$ as follows:

$$g(t) = \begin{cases} 0 & \text{for } 0 \le t \le z \\ -e^{-\lambda z}(t-z), & \text{for } t > z \end{cases}$$

Referring to any table of transform-function pairs, the Laplace transform of g(t) is $-\exp[-z(\lambda+s)]/s^2$. Thus the Laplace transform of H(t), i.e., the transform of $E_z[I(t)] + g(t)$, is

(3.3.1)
$$P_{H}(z,s) = \int_{0}^{\infty} e^{-st} H(t) dt = \frac{e^{-\omega(s)z}}{s\omega(s)} - \frac{e^{-z(\lambda+s)}}{s^{2}}.$$

Applying LINV to $P_H(z,s)$ we obtain an approximation of H(t). Using overscore to denote approximations, we compute the estimate of $E_Z[I(t)]$ as follows:

$$\bar{E}_{z}[I(t)] = \bar{H}(t) - g(t) = \bar{H}(t) + \begin{cases}
0 & \text{for } 0 \le t \le z \\
e^{-\lambda z}(t-z) & \text{for } t > z
\end{cases}$$

The bottom section of Figure 3.6 shows scaled $\overline{E}_z[I(t)]$ when the correction term, g(t), is included in the analysis. The results are considerably improved, and since the main tables will express $E_z[W(t)]$ in hundredths or thousandths, the slight inaccuracies which remain will not affect our tabulated results.

The correction term, g(t), is employed in our computations for both M/D/1 and M/E_k/1 queues. Equation (3.3.1) is incorporated into subprogram functions PMD1 and PMEK of the main programs. The gamma input process does not require the correction term because both $E_z[I(t)]$ and its derivative are continuous, i.e., the slope of $E_z[I(t)]$ is zero at t=z. The method is not applied to the Wiener process because process W(t) cannot be separated into four components, z+S(t)-t+I(t); thus, the third term of $P_W(z,s)$ cannot be interpreted as the Laplace transform of expected cumulative idleness in this case.

So far, our discussions in this section have covered some of the important details of the two main computer programs which produce the tables of scaled $\operatorname{E}_{Z}[W(t)]$ is Appendix C. Listings of these two main

programs, one for ρ = 1 and one for $\rho \neq 1$, are included in Appendix B. We now discuss the algorithm for M/E_k/1 queues with $\rho \neq 1$, and we briefly mention the differences in the algorithms for the other three processes and for ρ = 1. Figure 3.7 describes both the steps of the algorithm and the names of the subprograms which perform those steps.

The first program specifies the values of the parameters for subroutines LINV and ZERO. Values of ρ , scaled initial server load (z*), and sixteen scaled epochs (t*) are read from punched cards. The program performs computations for three pages of tables on each run. The three tables on a page have the same value of ρ and epochs t*, but each table has a different value of z*. For each table we compute $\overline{E}_z[W(t)]$ for the Wiener process at all sixteen epochs, then for the M/D/1 queue at the same sixteen epochs, followed by each of the seven M/E_L/1 queues and the gamma input process.

For each of the two processes the first step is conversion of z* and t* to unscaled values. As discussed in Section 2.3, the time scale conversion factor is $(1-\rho)^2/\sigma^2$; this factor is denoted ALFSQR in the subprograms RZWNR, RZMD1, RZMEK, and RZGAM. The value of σ^2 for the unscaled process is selected so that the formulas and numerical computations are subsequently simplified. The conversion factor for server load or wait is $|\mu|/\sigma^2$, denoted WTFCTR in the subprograms.

The initial trial value of $\omega(s)$ in the Newton-Raphson search for the solution of $\Phi[\omega(s)] = s$ is denoted OSTART. For a specific z^* and $\rho \leq 1$, OSTART is arbitrarily set equal to 1 for the first

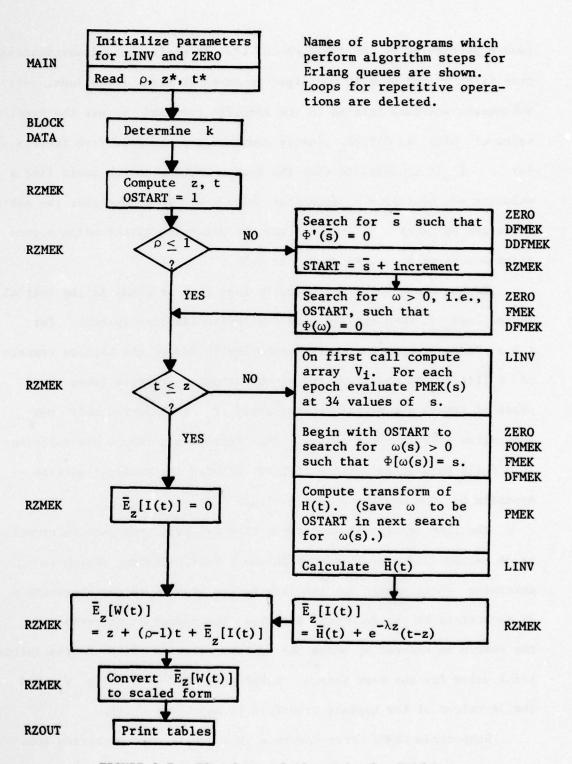


FIGURE 3.7. Flowchart of Algorithm for Tables

search pertaining to the first epoch t*. In preliminary investigations this first search usually converged in seven or eight iterations. All subsequent searches related to the specific z* and ρ use the previous value of $\omega(s)$ as OSTART, usually converging in four or five iterations. For $\rho > 1$ it is possible that the Newton-Raphson search could find a solution of $\Phi[\omega(s)] = s$ such that $\omega(s) < 0$. To ensure that the search converges to $\omega(s) > 0$ in this case, we determine OSTART using a two-stage search as described in Figure 3.7.

If the epoch to be evaluated is less than or equal to the initial server load z, then the expected cumulative idleness is zero. For t>z we use subprogram LINV to numerically invert the Laplace transform of $E_z[I(t)]$. On any given run of the computer programs (when three pages of tables are prepared), the array V_i is computed only when subprogram LINV is first called. This feature and others are explained in a fully documented version of LINV included in the first section of Appendix B, "Computer Program for Figure 3.2."

The last section of subprogram LINV evaluates the Laplace transform at 34 values. Each evaluation requires a Newton-Raphson search to determine $\omega(s)$, except for the case of the Wiener process where OMEGA is determined by the quadratic formula. The number of iterations in the search is reduced by using the current value of $\omega(s)$ as the initial trial value for the next search. Subprogram LINV uses array V_1 and the 34 values of the Laplace transform to calculate $\overline{H}(t)$.

Subprogram RZMEK first combines $\bar{H}(t)$ with the correction term to obtain $\bar{E}_z[I(t)]$, and then it computes $\bar{E}_z[W(t)]$ for the unscaled

process. In the final step $\tilde{\mathbf{E}}_{\mathbf{z}}[W(t)]$ is converted to scaled form and stored in the three-dimensional array SCLWT. When the computations for three tables (each with ten processes evaluated at sixteen epochs) are completed, subprogram RZOUT prints five copies of the tables with varying margins. Then the main program reads punched cards specifying ρ , z^* , and z^* for the next page of tables.

All computer programs listed in Appendix B were coded in IBM FORTRAN IV [IBM, 1971] and processed on an IBM 370/168 with the high-speed-multiply feature. The average computer time needed to evaluate scaled $\rm E_{Z}[W(t)]$ for one process at one epoch, i.e., a single value in the tables of Appendix C, was approximately 0.17 second.

CHAPTER 4

NUMERICAL RESULTS FOR EXPECTED SERVER LOAD

4.1. Using the Tables

In this section we explain how the tables in Appendix C can be used to obtain time-dependent mean server load in $M/E_{\rm k}/1$ queues. We will repeat only the notation and formulas from the previous chapters that are required for using the tables; the details of the underlying theory and numerical method will not be discussed here.

The system being studied is the standard $M/E_k/1$ queue. We assume that the analyst has specified the value of the mean arrival rate, denoted λ , and has determined that the service times (random variable S) can be described by a gamma or Erlang distribution. If an actual operating system is being studied and data are available, the analyst can use the methods described by Hora [1978] or Reinmuth [1971] to verify that the input process is Poisson. Similarly, a chi-square goodness-of-fit test can be used to compare the actual distribution of service times with tabulated values of the incomplete gamma function [Pearson, 1922]. In cases where the service times of an actual system do not exactly fit the gamma distribution, the analyst still may wish to use this theoretical distribution as an approximation.

We define the probability density function for this two-parameter gamma distribution using the mean, E(S), and the Erlang shape parameter, k (not restricted to integer values), as follows:

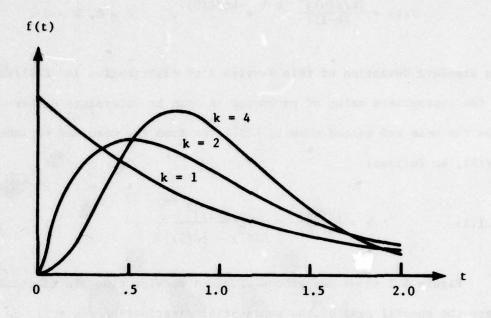
$$f(t) = \frac{[k/E(S)]^k}{(k-1)!} t^{k-1} e^{-kt/E(S)}, \quad t \ge 0, k > 0.$$

The standard deviation of this service time distribution is $E(S)/\sqrt{k}$, so the appropriate value of parameter k can be determined either from the mean and second moment, $E(S^2)$, or from the mean and variance, Var(S), as follows:

(4.1.1)
$$k = \frac{[E(S)]^2}{Var(S)} = \frac{[E(S)]^2}{E(S^2) - [E(S)]^2}.$$

Figure 4.1 shows two groups of gamma service time distributions, where the special case of the exponential distribution (k = 1) is included in both groups. The mean, E(S), of each distribution in the figure is equal to unity, so the standard deviation is $1/\sqrt{k}$ in each case. The top chart includes density functions with k = 2 and k = 4, i.e., with less variance than the exponential distribution. The limiting case as k becomes infinitely large is the degenerate distribution with constant service times, corresponding to the M/D/1 queue. The bottom chart of Figure 4.1 includes density functions with k = .5 and k = .2; these two distributions can be used to approximate systems where the service times have more variability than the exponential case.

The stochastic process being studied is server load, denoted W(t). This process is also called unfinished work, system backlog, or virtual waiting time, the latter because W(t) represents the time that an arrival at epoch t would wait before beginning service. The system may begin operation at epoch t=0 with some initial backlog, W(0)=z.



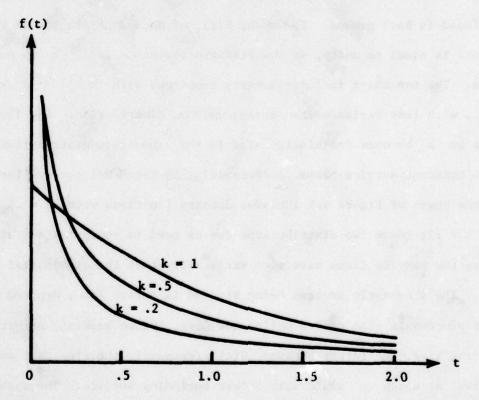


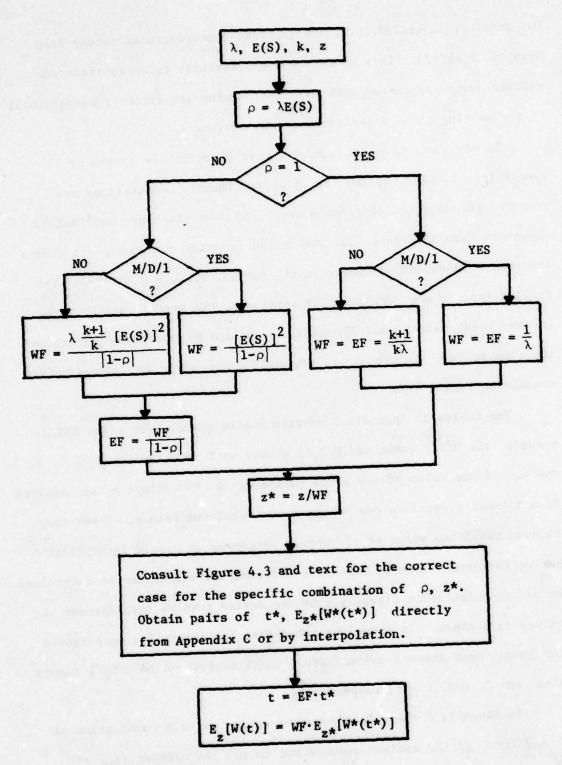
FIGURE 4.1. Erlang (Gamma) Density Functions with Mean = 1.

Our results, in scaled form, are for time-dependent mean server load, denoted $\mathbf{E}_{\mathbf{z}}[W(t)]$. This operating characteristic is appropriate when waiting costs associated with an actual system are directly proportional to the waiting times experienced by the customers.

To use the tables, the analyst must describe the system by specifying λ , E(S), k, and z. Some preliminary computations are required before consulting Appendix C, and these steps are outlined in flowchart form in Figure 4.2. We use an asterisk superscript to denote the scaled values which appear in the tables, e.g., $E_{z*}[W*(t*)]$. The scaling factors are WF, which is applied to the virtual waiting time (server load) values, and EF, which is applied to the epochs (points on the time scale). Values of ρ , z*, and k should be determined before consulting Figure 4.3.

The tables in Appendix C contain scaled results for eight M/G/1 queues: the M/D/1 queue and M/E $_k$ /1 queues with k = .2, .5, 1, 1.5, 2, 3, and 4. If the value of the shape parameter k determined by the analyst is not exactly equal to one of the seven tabulated values, either the closest tabulated value of k should be chosen or simple interpolation can be employed. In most cases the interpolated values can be determined by mental calculation directly from the scaled results for adjacent k values in a table. It should be noted that for interpolation purposes the gamma input process and the M/D/1 queue correspond to M/E $_k$ /1 queues with k = 0 and k = ∞ , respectively.

In Appendix C there is a separate table for each combination of ρ and z*. If the analyst chooses not to use the closest (ρ, z^*)



Note: EF = Epoch Factor for scaling
WF = Wait Factor for scaling

FIGURE 4.2. Flowchart Instructions for Using the Tables

THE RESIDENCE OF	z* = 0, .2, .4, .6, .8,1.0	0 < z* < 1; z* \neq .2,.4, .6,.8	z* = 2.,3., 4.	$1 < z^* < 4;$ $z^* \neq 2.,3.$	z* > 4.
ρ<.5	I	I I	I	I	I
ρ = .5,.6,.7, .8,.9	A	С	A	С	Н
.5 < ρ < .9; ρ ≠ .6,.7,.8	В	D D	В	D	Н
ρ = 1.0	A	С	A	С	G
ρ = 1.1,1.2 1.3,1.4 1.5,2.	A	С	F	F	F
1.1 < ρ < 2.; ρ ≠ 1.2,1.3 1.4,1.5	В	D	F	F	F
ρ > 2.0	E	E	F	F	F

FIGURE 4.3. Various Cases for (ρ, z^*) Combinations

combination which is tabulated, then interpolation between two or more tables will be required. Figure 4.3 illustrates the cases which might be encountered, and the appropriate methods for these cases are discussed below. For any specific (ρ, z^*) combination being studied, the objective is to determine pairs of t^* and $E_{z^*}[W^*(t^*)]$ to which the scaling factors, EF and WF, will be applied.

<u>Case A.</u> No interpolation is required. There is a table in Appendix C for this specific (ρ, z^*) combination, and the values for scaled epochs, t^* , and scaled mean server load, $E_{z^*}[W^*(t^*)]$, can be read directly from the table.

Case B. The exact value of z^* is tabulated, but the exact value of ρ falls between two tabulated values. First find the two tables that have the exact z^* with the two closest values of ρ . Then list the scaled epochs (t*) that appear on both tables, and also list both values of $E_{z^*}[W^*(t^*)]$ for each of the common epochs. For each epoch use simple interpolation to determine $E_{z^*}[W^*(t^*)]$ for the exact value of ρ .

<u>Case C.</u> The exact value of ρ is tabulated, but the exact value of z* falls between two tabulated values. The method is similar to Case B. Find the two tables with the exact ρ and the two closest values of z*; list the common values of t* with the two values of $E_{z*}[W*(t*)]$; and use simple interpolation. The initial behavior of

the system can also be determined as follows:

$$E_{z^*}[W^*(t^*)] = \begin{cases} z^* - t^* & \text{for } \rho < 1, \ 0 \le t^* \le (1-\rho)z^* \\ z^* & \text{for } \rho = 1, \ 0 \le t^* \le z^* \\ z^* + t^* & \text{for } \rho > 1, \ 0 \le t^* \le (\rho-1)z^* \end{cases}$$

Case D. The exact values of neither ρ nor z^* are tabulated, but the exact value of each does fall between two tabulated values. For one of the closest z^* values which are tabulated, perform the interpolation between the two closest values of ρ using the procedure of Case B. Use the same procedure for the other closest value of z^* . Finally, use the procedure of Case C to interpolate between those two sets of data. For example, if we are interested in results for $(\rho = .67, z^* = .44)$, we interpolate between $(\rho = .6, z^* = .4)$ and $(\rho = .7, z^* = .4)$ to obtain $(\rho = .67, z^* = .4)$. We also interpolate between $(\rho = .6, z^* = .6)$ and $(\rho = .67, z^* = .6)$ to obtain $(\rho = .67, z^* = .6)$. Finally, we interpolate between $(\rho = .67, z^* = .4)$ and $(\rho = .67, z^* = .6)$ to obtain results for $(\rho = .67, z^* = .4)$.

Case E. For situations where $\rho > 2$ and $0 < z^* < 1$, the Wiener process provides an upper bound for $E_{z^*}[W^*(t^*)]$ in any queueing system. Because of the scaling used in the tables, the tabulated values of $E_{z^*}[W^*(t^*)]$ for the Wiener process depend only on z^* , not on ρ . Therefore, any table with $\rho \ge 1.1$ and the specified value of z^* can be used to determine the upper bound. Interpolation can be used if the

exact value of z^* falls between the tabulated values. In these same situations a lower bound for $E_{z^*}[W^*(t^*)]$ is $z^* + t^*$.

<u>Case F.</u> For situations with both $\rho > 1$ and $z^* > 1$, a very accurate approximation of $E_{z^*}[W^*(t^*)]$ is $z^* + t^*$. The accuracy is best when both ρ and z^* are large, but even for $\rho = 1.1$ and $z^* = 1$ the maximum error (M/D/1 queue, $t^* = 1.5$) is only 1.4 percent.

Case G. For $\rho=1$ and $z^*>4$, the initial behavior is $E_{z^*}[W^*(t^*)] = z^* \quad \text{for} \quad 0 \le t^* \le z^*. \quad \text{For} \quad t^*>z^*, \text{ the results for}$ $z^*=4$ provide a loose lower bound.

<u>Case H.</u> For $.5 \le \rho \le .9$ and $z^* > 4$, the initial behavior is $E_{z^*}[W^*(t^*)] = z^* - t^*$, for $0 \le t^* \le (1-\rho)z^*$. For $t^* > (1-\rho)z^*$, the results for $z^* = 4$ provide a lower bound.

<u>Case I.</u> For systems with $\rho < .5$ the approach to equilibrium is very slow. If the initial server load $z^* > 0$, the initial behavior is the same as Case H. For $z^* \le 4$ the results for the same queue (i.e., the same value of k) with $\rho = .5$ provide an upper bound.

Other Cases. For situations with $.9 < \rho < 1.0$ and $0 \le z^* \le 4$, the tabulated results for the Wiener process with any $\rho < 1$ and the appropriate z^* provide a tight upper bound for any queueing system. Likewise, for situations with $1.0 < \rho < 1.1$ and $0 \le z^* \le 4$, the

tabulated results for the Wiener process with any $\, \rho > 1 \,$ and the appropriate _z* provide a tight upper bound. In both cases the Wiener approximation is most accurate for values of $\, \rho \,$ closest to unity.

The tables in Appendix C can be used along with Figures 4.2 and 4.3 to obtain explicit results for the transient behavior of mean server load in a wide range of M/G/1 queues. In many applications the analyst will perform some sensitivity analysis by using various estimates of λ , E(S), k, and z and comparing the results. In cases where $\rho < 1$ the analyst may be interested only in the speed with which the system approaches steady-state. Expressing the Pollaczek-Khintchine formula with our notation, the mean of the steady-state distribution for server load (often called expected waiting time in the queue, excluding service) is

$$E_{\mathbf{z}}[W(\infty)] = WF/2$$
, $\mathbf{z} \geq 0$, $\rho < 1$.

4.2. Several Sample Problems

In this section we illustrate the methods described in Section 4.1 by working several sample problems. The first three problems show how changes in the variance of service times affect the transient behavior of mean server load and the steady-state value, assuming the other parameters of the system are held constant. The fourth and fifth problems illustrate using interpolation to determine results for queues when the exact value of ρ or z^* is not tabulated. The determination of waiting

costs using time-dependent mean server load is discussed and applied to one of the sample problems.

The following problems are adapted from an example described by Hillier and Lieberman [1974, p. 436]. A university plans to lease a small batch-processing computer for its students to use. It is expected that the students will submit programs to be run every 3 minutes on the average and that the times between submission of programs have an exponential distribution. The computer under consideration could process an average of 25 typical student programs per hour if it were run continuously. We will examine the transient behavior of this M/G/1 system for an eight-hour period, assuming that the system begins operation with no backlog. Thus far we have specified λ , E(S), and z. For Problem A we assume that the processing times for the students' programs are exponentially distributed, i.e., k = 1. Referring to Figure 4.2, we perform the initial computations, using hours as the time unit.

Problem A.

$$\lambda = 20$$
, E(S) = 1/25, k = 1, z = 0
$$\rho = 20 \cdot (1/25) = .8$$

$$WF = \frac{20 \cdot 2 \cdot (1/25)^2}{1 - .8} = .32$$

$$EF = .32/.2 = 1.6$$

$$z* = 0/.32 = 0$$

Referring to Figure 4.3 with ρ = .8 and z* = 0, we see that Case A is appropriate, i.e., the exact values of ρ and z* appear in a table in Appendix C. Referring to the appropriate table, we note that the sixteen scaled epochs t* are between .02 and 4. Eight points should be sufficient to describe the queue's behavior; the selected values of t* and $E_{z*}[W*(t*)]$ from Appendix C are shown below in the first two columns. Referring to Figure 4.2, we compute t and $E_{z}[W(t)]$ by multiplying the scaled values times the scaling factors. The desired results are shown below in the last two columns.

Problem A.

t*	Ez*[W*(t*)]	t	E _z [W(t)]
.1	.164	.16	.052
.2	.230	.32	.074
.4	. 306	.64	.098
.6	.351	.96	.112
1.	.405	1.6	.130
2.	.461	3.2	.148
3.	.482	4.8	.154
4.	.491	6.4	.157

The steady-state mean server load in Problem A is WF/2 = .16 hour (9.6 minutes), or four average service times. After approximately one hour of operations (t = .96), mean server load (.112 hour) in this M/M/1 queue is about 70% of steady-state; after three hours it achieves

90% of the steady-state value. For queues with $\rho < 1$, $E_{Z^*}[W^*(t^*)]$ approaches a steady-state value of .5; therefore, the percentage of steady-state achieved at any epoch can be determined simply by doubling the $E_{Z^*}[W^*(t^*)]$ value.

For Problem B we consider a system with the same arrival rate and the same mean service time, but we assume that the service times are constant (deterministic). In the context of our original computer-leasing problem, it might be that the student programs are actually uniform or that this specific computer has very little variation in processing times for the programs being run. Referring to Figure 4.2, we perform the computations for this M/D/1 queue.

Problem B.

$$\lambda = 20$$
, E(S) = 1/25, k = ∞ , z = 0
$$\rho = 20 \cdot (1/25) = .8$$

$$WF = \frac{20 \cdot (1/25)^2}{1 - .8} = .16$$

$$EF = .16/.2 = .8$$

$$z* = 0/.16 = 0$$

Referring to Figure 4.3, we see that the exact values of ρ and z^* are tabulated, and we select eight points from the appropriate table in Appendix C as shown below. The tabulated values are mutiplied by EF and WF to obtain t and $E_z[W(t)]$, respectively.

Problem B.

t*	Ez*[W*(t*)]	t	E _z [W(t)]	
.2	.242	.16	.039	
.4	.316	.32	.051	
.6	. 360	.48	.058	
1.	.412	.8	.066	
1.5	.446	1.2	.071	
2.	.465	1.6	.074	
3.	.484	2.4	.077	
4.	.492	3.2	.079	

The steady-state mean server load in Problem B is WF/2 = .08 hour (4.8 minutes), or two service times. We observe that after a half-hour of operation (t = .48) mean server load (.058 hour) is about 70% of steady-state; after one and a half hours (t = 1.6), it has reached 90% of the steady-state value (.074 relative to .08). Compared with the M/M/1 queue of Problem A, this M/D/1 queue has a lower steady-state mean server load and it approaches that equilibrium value much faster.

For Problem C we consider a system with the same prameters as Problems A and B, except that the service times have wide variability. We assume that the mean service time is 2.4 minutes (1/25 hour), as before, but that the standard deviation of the service times is 5.4 minutes. Referring to equation (4.1.1), the appropriate Erlang shape

parameter is k = .2. The computations from Figure 4.2 and the selected points from Appendix C are shown below.

Problem C.

$$\lambda = 20$$
, E(S) = 1/25, k = .2, z = 0
$$\rho = 20 \cdot (1/25) = .8$$

$$WF = \frac{20 \cdot 6 \cdot (1/25)^2}{1 - .8} = .96$$

$$EF = .96/.2 = 4.8$$

$$z* = 0/.96 = 0$$

t*	E _{z*} [W*(t*)]	t	E _z [W(t)]	
.1	.156	.48	.150	
.2	.223	.96	.214	
.4	.299	1.92	.287	
.6	. 345	2.88	.331	
.8	.377	3.84	. 362	
1.	.400	4.8	. 384	
1.5	.438	7.2	.420	
2.	.459	9.6	.441	

The steady-state mean server load in Problem C is WF/2 = .48 hour (28.8 minutes), or twelve average service times, which is much higher than the M/M/1 and M/D/1 cases. We observe that about 70% of steady-state is achieved after three hours and that approximately eight hours of operation are needed to reach 90% of the steady-state value.

Figure 4.4 shows mean server load curves for Problems A, B, and C, where these three queues have the same values of λ and E(S). We observe that the queues with higher variance of service times, i.e., with lower values of k, have a higher steady-state mean server load. Furthermore, the approach to equilibrium is slower for the queues with the higher steady-state value. We also note from Figure 4.4 that the ordering of the curves is the opposite of the ordering of the scaled values in the table (ρ = .8, z* ≈ 0) in Appendix C; this reordering is explained by the different scaling factors, WF and EF, for the three queues.

Returning to our computer-leasing problem, we next consider a computer that can process an average of 30 programs per hour, i.e., E(S) = 1/30. We assume that the processing times are exponentially distributed, i.e., k = 1, and that the arrival pattern is the same as Problems A, B, and C. For Problem D we will assume that students can leave programs during the night for processing when the facility opens. It is estimated that "about a dozen" programs will arrive during a typical night; thus, the initial server load (assuming 12 jobs at 1/30 hour each) is .4 hour. Referring to Figure 4.2, we perform the following calculations.

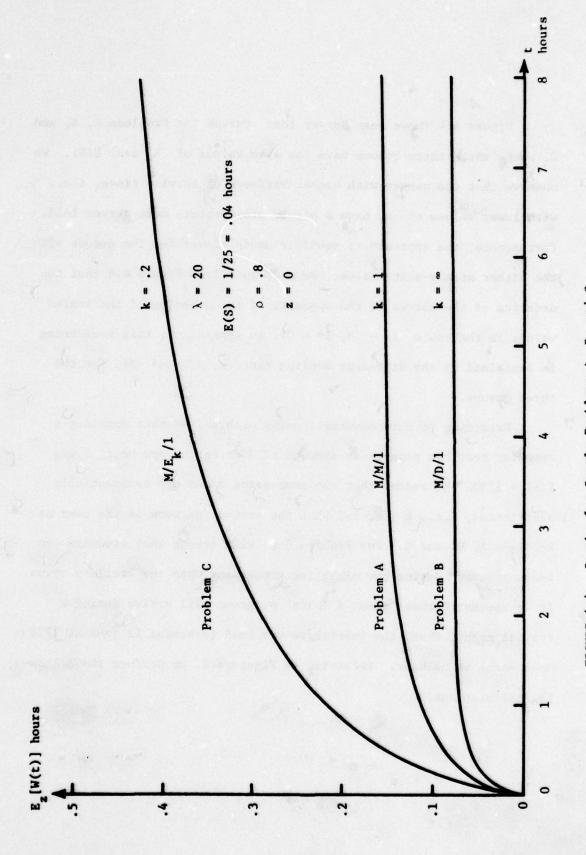


FIGURE 4.4. Graphs for Sample Problems A, B, and C.

Problem D.

Values from Appendix C

			Interpolated		
	E_* [W*(t*	;)], z*=3	E _{z*} [W*(t*)]		E _z [W(t)]
t*	ρ = .6	ρ = .7	ρ = .67	t	ρ = .67
.5	2.5	2.5	2.5	.2	. 333
1.	2.0	2.0	2.0	.4	.267
1.5	1.528	1.551	1.543	.6	.206
2.	1.195	1.225	1.215	.8	.162
2.5	.978	1.004	.995	1.	.133
3.	.835	.854	.848	1.2	.113
4.	.673	.681	.678	1.6	.090
5.	.593	.596	.595	2.	.079
6.	.552	.552	.552	2.4	.074
7.	.530	.529	.529	2.8	.071
8.	.517	.517	.517	3.2	.069
10.	.506	.506	.506	4.	.067
12.	.502	.502	.502	4.8	.067

Problem D.

$$\lambda = 20$$
, E(S) = 1/30, k = 1, z = .4
 $\rho = 20 \cdot (1/30) = .67$
WF = $\frac{20 \cdot 2 \cdot (1/30)^2}{1 - .67} = .133$
EF = .133/.333 = .4
z* = .4/.133 = 3

Referring to Figure 4.3 with ρ = .67 and z^* = 3, we see that Case B is appropriate. The two relevant tables are $(\rho$ = .6, z^* = 3) and $(\rho$ = .7, z^* = 3), and we list the values of t^* appearing on both tables. Those thirteen values and the corresponding values of $E_{z^*}[W^*(t^*)]$ from the two tables are shown in the first three columns below. For each scaled epoch we interpolate to find the approximate value that is two-thirds (or about seven-tenths) of the difference between the values for ρ = .6 and ρ = .7. The interpolated value is shown below in the fourth column. Finally, we obtain t and $E_z[W(t)]$ by multiplying t^* and the interpolated values of $E_{z^*}[W^*(t^*)]$ times EF and WF, respectively.

The steady-state mean server load in Problem D is WF/2 = .067 hour (4 minutes), or two mean service times. We observe that transient mean server load is within 10% of the steady-state value after approximately two and a half hours of operation. Figure 4.5 shows the curve for Problem D and the curve for the same queue with no initial load. (The calculations for the latter case are not shown here.) The system with z = 0 reaches 90% of steady-state after only .8 hour of operation.

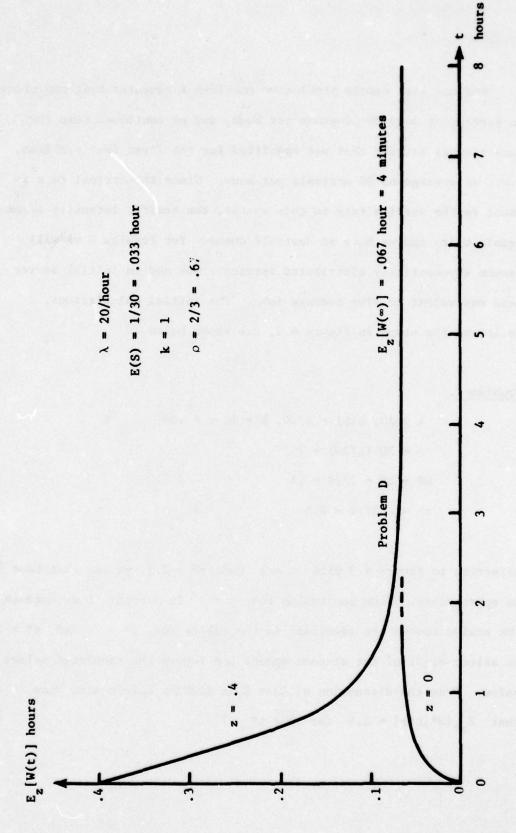


FIGURE 4.5. Graph for Sample Problem D.

For our last sample problem we consider a computer that can process an average of only 20 programs per hour, and we continue to use the same arrival pattern that was specified for the first four problems, i.e., an average of 20 arrivals per hour. Since the arrival rate is equal to the service rate in this system, the traffic intensity parameter equals unity and we have an unstable queue. For Problem E we will assume exponentially distributed service times and an initial server load equivalent to five average jobs. The initial calculations, following the steps in Figure 4.2, are shown below.

Problem E.

$$\lambda = 20$$
, E(S) = 1/20, k = 1, z = .25
 $\rho = 20 \cdot (1/20) = 1$
WF = EF = 2/20 = .1
 $z* = .25/.1 = 2.5$

Referring to Figure 4.3 with $\rho=1$ and $z^*=2.5$, we see that Case C is appropriate. From the tables for $\rho=1$ in Appendix C we note that the scaled epochs are identical in the tables for $z^*=2$ and $z^*=3$. We select eight of the sixteen epochs and record the tabulated values below. From the discussion of Case C in Section 4.1 we also know that $E_{z^*}[W^*(t^*)]=2.5$ for $0 \le t^* \le 2.5$.

Problem E.

Values fro	om Appendix	: C	Interpolated		
	E _{z*} [W*(t*	$[, \rho = 1]$	E _{z*} [W*(t*)]		E _z [W(t)]
t*	z* = 2	z* = 3	z* = 2.5	t	z = .25
0 to 2.5	with some in		2.5	0 to .25	.25
3	2.056	3.000	2.528	.3	.253
4	2.149	3.016	2.583	.4	.258
6	2.362	3.108	2.735	.6	.274
8	2.577	3.238	2.908	.8	.291
10	2.784	3.383	3.084	1.	. 308
20	3.680	4.115	3.898	2.	. 390
40	5.052	5.364	5.208	4.	.521
80	7.068	7.289	7.179	8.	.718

If the M/M/l system described in Problem E continues operating indefinitely, the mean server load will increase without bound. In many actual systems the arrival rate equals or exceeds the service rate for short periods of time, but such situations may be acceptable if the cost of providing faster service exceeds the cost associated with the high waiting times.

The five sample problems show that the methods described in this chapter can be used to determine time-dependent mean virtual waiting time, an important operating characteristic of M/G/1 queues. In actual applications the analyst may wish to determine the cost

associated with $E_{\rm Z}[W(t)]$ when analyzing or designing a system. In their discussion of decision models for queueing systems, Hillier and Lieberman [1974, Ch. 10] discuss appropriate cost functions for various problems and apply their methods to steady-state distributions of both the number of customers in the system and the waiting times of individual customers. As previously mentioned, our results for expected server load are appropriate for evaluating costs only in systems where the waiting costs are directly proportional to waiting times. We next describe a method for evaluating the waiting cost in situations where the transient behavior of the system is important.

Let c be the cost per unit of waiting time experienced by an individual customer. The cost associated with a customer arriving at epoch t is c·W(t), where there may be some initial server load W(0) = z for process $W(\cdot)$. For any specific epoch t, W(t) is a random variable; in our situation the cost function is linear, and the expected cost for a customer arriving at epoch t is c·E_z[W(t)]. In cases where the cost function is non-linear, we would need more information about the distribution of W(t), not just its mean, in order to determine the expected cost of a customer arriving at epoch t. The probability that a customer will arrive during interval (t, t+ Δ t) is Δ dt, and such a customer would incur expected cost c·E_z[W(t)]. Thus, the expected waiting cost during a period from t = 0 to t = T is

(4.2.1)
$$\int_{0}^{T} c E_{z}[W(t)] \lambda dt = c\lambda \int_{0}^{T} E_{z}[W(t)] dt .$$

Note that the integral on the right-hand-side of (4.21) is simply the area under the curve for transient mean server load.

For a queueing system that approaches steady-state quite slowly, the analyst should use the integral expression, not the equilibrium mean waiting time, to evaluate expected cost. To illustrate the difference, consider again the system described in Problem C. Assume the university has determined that for this situation the appropriate cost associated with making a student wait for a program to be processed is \$10 per hour. If we approximate the mean waiting time using the steady-state value of .48 hours, then the expected waiting cost for an individual student arriving with a computer program to be run is \$4.80. The mean arrival rate is 20 per hour, so the expected waiting cost is \$96 for each hour the system is in operation. Thus, during an eighthour day the total expected waiting cost is \$768. This total cost figure is based on the equilibrium mean waiting time, but from Figure 4.4 we can see that $E_z[W(t)]$ for Problem C is substantially less than the equilibrium value.

For a more accurate determination of total waiting cost during the eight-hour period, we can evaluate the integral on the right-handside of equation (4.2.1) graphically by plotting the curve from Figure 4.4 on a fine grid and counting the squares. Using this method the value of the area under the curve is approximately 2.66 hours². (The corresponding value using the equilibrium mean server load is .48.8 = 3.84 hours².) Multiplying the area under the curve by $c \cdot \lambda$, we find that the correct total expected waiting cost for the eight-hour period is \$10.20.2.66 = \$532.

In this particular problem the use of the equilibrium value produces a total cost figure approximately 44% greater than the correct amount. The example illustrates the importance of using time-dependent results when the approach to equilibrium is slow. The graphical technique can also be used to determine expected waiting cost for unstable queueing systems $(\rho \geq 1)$, where all behavior is transient and no equilibrium is ever achieved.

CHAPTER 5

CONCLUSIONS

The objective of this research was to provide tables for time-dependent mean server load in M/G/1 queueing systems. Chapter 2 presented the theoretical framework for the server load process, the Laplace transform of $E_Z[W(t)]$, and a scaling procedure that allowed us to reduce the number of parameters required to describe a specific system from four $[\lambda, E(S), z, k]$ to three $[\rho, z^*, k]$. In Chapter 3 we considered a technique for inverting the Laplace transform, and we conducted extensive checks and comparisons to ensure the accuracy of our numerical results. The sample problems in Chapter 4 illustrated that a practitioner can easily use our tabulated results in Appendix C by following simple step-by-step procedures. We now discuss some additional results of this research, including a study of the error associated with the Brownian approximation, and we offer some suggestions for future research.

The scaling procedure employed in this research produces curves for $E_{z*}[W*(t*)]$ that are monotonically ordered by each of the three parameters $(\rho, z*, \text{ and } k)$, and these consistent orderings facilitated the interpolation procedures described in Chapter 4. Referring to any single table in Appendix C, i.e., fixing the values of ρ and z*, we observe that the curves for $E_{z*}[W*(t*)]$ increase monotonically as k, the shape parameter of the Erlang service time distribution, increases. (This ordering of the transient curves by the shape parameter has not been proven analytically and is a possible topic for future theoretical

research.) As we would expect, the upper bound (corresponding to $k = \infty$) is the curve for the M/D/1 queue. But our numerical results also indicate that a lower bound can be identified: the gamma input process (corresponding to k = 0). We can think of these upper and lower bounds as defining an "envelope", and we might hypothesize that this envelope contains the curves, in scaled form, for all M/G/1 queueing systems, not just for the M/E_k/1 queues that are studied here. Figure 5.1 shows the upper and lower bounds for the $(\rho = .5, z* = 0)$ case. We observe that the envelope is relatively narrow; specifically, its maximum width is 11% of the steady-state value in this case. For z* = 0 and $\rho = .6, .7, .8$, and .9, the maximum width is 9%, 7%, 5%, and 3%, respectively. We could use these bounds to obtain approximate results for any M/G/1 queueing system for which λ , E(S), E(S²), and z are known, without specifying the exact shape of the service time distribution.

Another monotonic ordering is observed by fixing the values of k and z* and varying ρ . For example, if we examine the curves for the M/D/1 queue (k = ∞) with z* = 0, we observe that for $\rho < 1$ the scaled curves increase monotonically as ρ increases. (In Appendix C this comparison requires examining tables that are each three pages apart.) The M/D/1 curves seem to be approaching the curve for the Wiener process as ρ increases. This observation has not been proven analytically, but the heavy traffic (or Brownian motion) approximation for the equilibrium distribution of M/G/1 server load is derived by examining, the scaled form, the limit of a sequence of queues as ρ

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t*							
ρ	.1	.2	.4	.6	1.	2.	4.
2.0	91	60	39	30	21	13	7
1.5	56	38	25	19	14	8	5
1.4	46	32	21	16	12	7	4
1.3	37	25	17	13	10	6	3
1.2	25	18	12	9	7	4	2
1.1	13	9	6	5	4	2	1
.9	14	10	6	4	3	1	0
.8	35	23	14	10	7	3	1
.7	65	42	26	18	12	6	2
.6	108	68	41	30	19	9	3
.5	171	109	67	48	31	15	5

The tabulated error is the difference between the Wiener and the gamma input values, divided by gamma input value, and rounded to the nearest whole per cent.

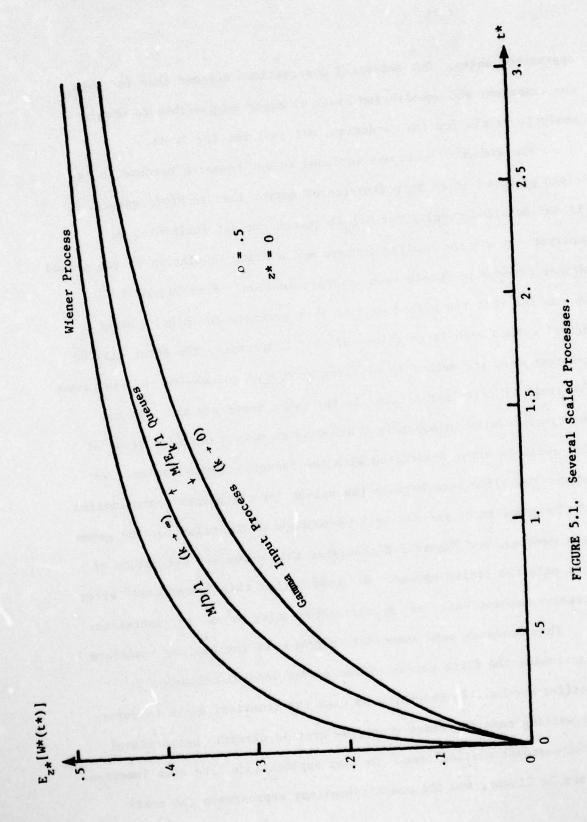
FIGURE 5.2. Percentage Error of the Wiener Approximation, $z^* = 0$

relationship using a quadratic function. In such cases knowledge of both the first and second moments of the server load distribution is required. Prabhu [1965] derived the double Laplace transform for the time-dependent server load distribution in M/G/1 queues. If we evaluate the second derivative of that expression at zero, we obtain the Laplace transform for the time-dependent second moment of the distribution. The LINV algorithm can then obtain the second moment at specified epochs for use in a quadratic cost function. There may be other queueing applications for this relatively efficient and accurate numerical technique for Laplace transform inversion.

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approaches unity. Our empirical observations suggest that for both the transient and equilibrium cases it might be possible to obtain analytic proofs for the ordering, not just for the limit.

The Wiener process was included in our research because it is often proposed as an approximation of server load in M/G/1 queues. If our detailed results for $M/E_k/1$ queues are not available, an analyst can use the scaling factors and a brief tabulation of the scaled Wiener process to obtain such an approximation. From Figure 5.1 we can see that the method will be most accurate for M/D/1 queues or $M/E_{\rm b}/1$ queues with large values of k. Conversely, the error will be greatest when the method is used for an $M/E_k/1$ queue with k near zero. The limit of this "worst case" is the gamma input process, and our numerical results in Appendix C allow us to make a detailed study of the possible error associated with the Brownian approximation. We express the difference between the values for the Wiener approximation and the gamma input process as a percentage of the value for the gamma input process, and Figure 5.2 tabulates this error for all values of p at selected scaled epochs. We observe that this "worst case" error decreases monotonically as p approaches unity or as t* increases.

This research used numerical inversion of the Laplace transform to determine the first moment of the server load distribution at specified epochs. In Chapter 4 we used the transient curve to determine waiting cost for cases where the cost is directly proportional to the customer waiting time. In many applications, the cost function may not be linear, and the practitioner may approximate the cost

APPENDIX A

SUMMARY OF NOTATION

This list includes brief definitions of the notation used most frequently in this dissertation.

A(t)	stationary Poisson process, the M/G/1 arrival process
c	cost per unit of waiting time
E(·)	expectation, first moment, of a random variable
E _z [·]	expected value, conditional on initial load z
EF	scaling factor for epochs
Fa	an approximation based on Stehfest's algorithm
F(·)	cumulative probability distribution for S
F*(s)	Laplace-Stieltjes transform of F(·)
f(')	probability density function
f _n (t; a)	observational density function with parameters a and n
G(t)	gamma process
g(t)	correction term function
H(t)	$E_{z}[I(t)] + g(t)$
I(t)	cumulative idleness process
k	Erlang shape parameter
N	number of values of the Laplace transform used by Stehfest's algorithm to obtain an approximation
P(t)	any function of interest
\bar{P}_n	an approximation of $P(t)$ based on n values of the Laplace transform using Gaver's method

- $P_{H}(z,s)$ Laplace transform of H(t), conditional on z
- $P_{W}(z,s)$ Laplace transform of W(t), conditional on z
- p(s) any Laplace transform function
- S, S, service time random variable
- S(t) compound Poisson process, an input process
- T random variable for an observational epoch
- Var[·] variance of a random variable or process
- W random variable for equilibrium server load
- WF scaling factor for waiting time
- W(t) server load, or virtual waiting time, process
- X(t) net input process, a Levy process
- z deterministic initial server load, W(0)
- a_i error components of an asymptotic expansion
- α , β shape and scale parameters for gamma density function
- Γ(·) gamma function
- λ mean arrival rate, parameter for A(t)
- μ -E[X(t)]/t
- ξ(t) standard Wiener process
- ρ traffic intensity parameter
- $\sigma^2 \qquad Var[X(t)]/t$
- Φ(·) exponent function
- $\omega(s)$ solution of functional equation $\Phi[\omega(s)] = s$
- overscore, or "bar", denotes approximation
- * asterisk superscript, denotes a scaled value

APPENDIX B

LISTINGS OF THE COMPUTER PROGRAMS

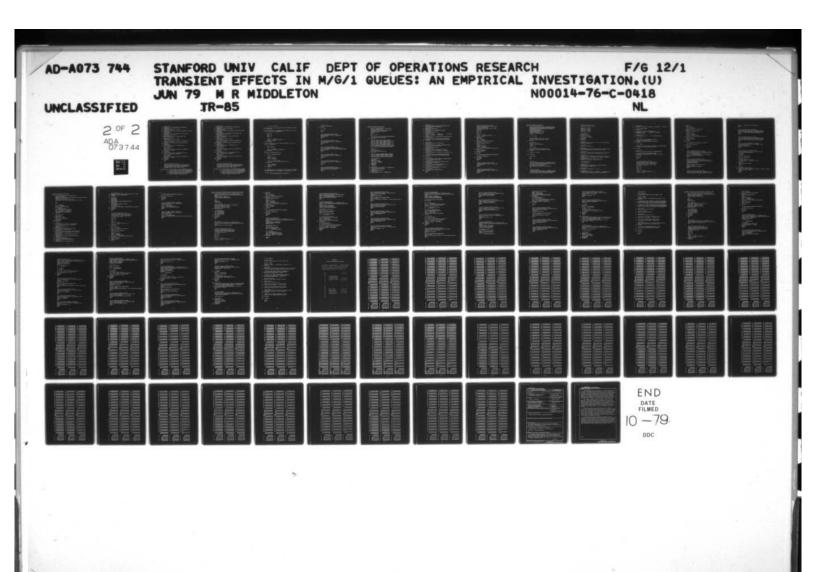
- 1. Figure 3.2, including the documented version of LINV
- 2. Figure 3.4
- 3. Figure 3.5
- 4. Figure 3.6
- 5. Appendix Tables, $\rho \neq 1$
- 6. Appendix Tables, $\rho = 1$

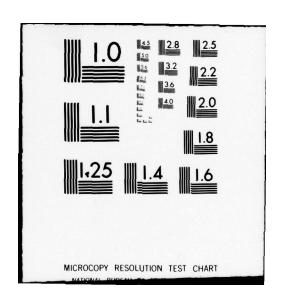
COMPUTER PROGRAM FOR FIGURE 3.2

```
IMPLICIT REAL #8 (A-H, O-Z), INTEGER (I-N)
    COENCE/LAFLAC/EPUCH(10), APPROX(10,6)
    LINEASION EXACT (10,6), STEHF (10,6), V (50)
    EXTEBBAL P1, F2, P3, P4, F5, P6
    M=0
    N = 34
    LATA STEHF /. 56555,. 39912,. 32655,. 28278,
   1.25174, .22989, .21322, . 19956, . 18814, . 17796,
   1. 16566, 1. 32543, 4. 4, 354,
   110.60342,20.70645,35.76832,56.82535,84.82735.
   1120.78473,165.66749,
   1.98775, .91001, .63826, .30966,
   1-.02119,-.31527,-.57254,-.76869,-.91049,-.98949,
   1-.57762,-1.27084,-1.67544,
   1-1.90352,-2.18727,-2.36870,-2.52276,-2.05740,
   1-2.77390, -2.88091,
   1. 36798, . 13557, . 05043, . 01849,
   1.00640,.00195,.00036,-.00006,-.00047,-.00020,
   1-.66533,-.32531, 1.02575, 2.39533,
   12.78644, 1.21692, -3.32956, -11.82953, -25.283 > 3, -44.86511/
    DO 30 I=1,10
    EFOCH(I)=I
    1 = I
    EXACT (I, 1) = 1/DSQAT (1+3.1415926535897932384626433)
    EXACT (I,2) =T+T+T/6
    EXACT (I,3) = DSIN (DSUNT (2*T))
    EXACT(I,4) = -DLOG(T) - G.57721560490153386061
    EXACT (I,5) = DEXP(-T)
    EXACT (1,6) = 1- 3+T+ 3+T+T/2-T+T+T/6
    CALL LINV(£1,N,T,AFPROX(I,1),V,H)
    CALL LINV (P2, N, T, APPAOX (I, 2), V,M)
    CALL LINV (P3, N,T, APPROX (I,3), V,M)
    CALL LINV (P4,N,T,APPROX (I,4),V,K)
    CALL LINV (F5, N, T, APPROX (I, 5), V, A)
    CALL LIAV (16, N.T. APPHOX(I,6), V, d)
30
    CONTINUE
    DU 80 N COPY=1,5
    WAITE (0,4)
   FORMAT (1H1,////)
    MARGIN = 4 + NCOPY
    LO & IMRGN = 5, MARGIN
    WAITE (D, O)
   FORMAL (1H )
    CUNTINUE
    WRITE (0,40)
   PORMAT(1H , 35%, FIGURE 3.2 INVERSION OF TEST',
   1' FU MCLICAS')
    WA ITE (6,42)
42 FORMAT (1H , 35 X, 39 ('-'))
    .EITE (6,43)
    FORMAT (1HC, 33x, 2 ('APPROXIMATE P(T) USING', 21x))
```

COME UTER PAUGRAM FOR FIGURE 3.2

```
IMPLICIT REAL +8 (A-H, 0-2), INTEGER (I-N)
    COEMUN/LAFLAC/EPUCH(10), APPRUX(10,6)
    LIMEASIUM EXACT (10,6), STEHF (10,6), V (50)
    EXTEBNAL P1, F2, P3, P4, F5, P6
    M=0
    N = 34
    LATA STEHF /. 56555,. 39912,. 32655,. 28278,
   1. 25174, .22989, .21322, . 19956, . 18814, . 17796,
   1. 16566, 1. 32543, 4. 4, 354,
   110.60342,20.70845,35.78832,56.82535,84.82735,
   1120.78473,165.66749,
   1.98775, .91001, .63826, .30966,
   1-.02119,-.31527,-.57254, -.76869,-.91049,-.98949,
   1-.57762,-1.27084,-1.67544,
   1-1.90352,-2.18727,-2.36870,-2.52276,-2.05740,
   1-2.77390, -2.88091,
   1. 36798, . 13557, . 05043, . 01849,
   1.00640,.00195,.00036,-.00006,-.00047,-.00020,
   1-.66533,-.32531, 1.02575, 2.39533,
   12.78644,1.21692,-3.32956,-11.82953,-25.28333,-44.86511/
    LO 30 I=1,10
    EPOCH(I)=I
    1 = I
    EXACT (I, 1) = 1/DSQLT (1+3.1415926535897832384626433)
    EXACT (I,2) =T+T+T/6
    EXACT (I,3) = DSIN (DSLAT (2*T))
    EXACL(I,4) = -DLOG(L) - G.57721560490153386061
    EXACT (I,5) = DEXP(-T)
    EXACT (1,6) = 1-3+T+3+T+T/2-T+T+T/6
    CALL LINV ( 1 1, N ,T , APPROX (I, 1) , V , M)
    CALL LINV (P2, N, T, APPAOX (1, 2), V,M)
    CALL LINV (P3, N,T, APPROX (I,3), V,M)
    CALL LINV (P4,N,T,APPROX (I,4), V,M)
    CALL LINV (F5, N, T, APPROX (I, 5), V, M)
    CALL LIAV (16, N.T. APPHOX(I,6), V,M)
    CONTINUE
    DC 80 MCOPY=1,5
    WAITE (0,4)
   FORMAT (1H1,////)
    MARGIN = 4 + NCOPY
    LO & IMRGN = 5, MARGIN
    WAITE (0,0)
   FORMAL (1H )
    CUNTINUE
    WRITE (0,40)
    FORMAT (1H , 35%, FIGURE 3.2 INVERSION OF TEST',
   1' FUNCAIGNS')
    WA ITE (6,42)
42 FORMAT (1H , 35 X, 39 ('-'))
    .EITE (6,43)
    FORMAT (1HC, 33X, 2 ('APPROXIMATE P(T) USING', 21X))
```





```
mkITE (6,44)
      FORMAT (1H , 35%, 2 (17HSTEHPEST'S METHOD, 20%))
      WRITE (6,45)
      FORMAT (1H , 32X,2 (25 ('-'), 18X))
  45
      WKITE (6,46) N.N
     FORM AT (1H , 11X, 'T', 6X, 2('EXACT F(T)', 9X, 'N=', 12,9X,
     1'N=10', 7X))
      WEITL (6,48)
     FORMAT(1H ,10x,'---',2(2x,14('-')),2x,9('-'),
12(2x,14('-')),2x,9('-'))
      WEITE (0,50)
      FURMAD (1HO, 20X, 'F (T) = 1/SQAT (PI+T) ,24X,
     1'F(T) = -C-Lh(T)'
      WE IT & (6,52)
  52 FORMAT (14 , 16x , 2 (41 ('-') , 2X))
      DU 50 I=1,10
      WEITE (0,54) EPOCH (I), (EXACT (I,J), AFFECX (I,J),
     152EnF(I,0), J=1,4,3)
     FORMAT (1H , 10X,F4. 1, 2(2(2X,F14.10),2X,F9.5)).
  56 CONTINUE
      WRITE (6,60)
  60 FORMAT (1HO, 26X, 'F (T) = (T*T*T) /6',27X,
     1'F(2) = EXF(-T)'
      WRITE (6,52)
      ₩0 66 I=1,10
      WRITE(6,54) EPOCH(I), (EXACT(I,J), APPROX(I,J),
     15TEHF (1,J), J=2,5,3)
  66 CONTINUE
      WE ITE (0,70)
  70 FOBBAT (1HO, 26X, F(T) = SIN (SURT (2+T)) 1,22%,
     1 °E (T) = 1-3+1+3+1+1/2-1+T+T/6 °)
      WEITE (0,52)
      DO 76 I=1,10
      WAITE (6,54) EPOCH (I), (EXACT (I,J), APERGX (I,J),
     1STEHF (1,J), J=3,6,3)
      CUNT INUL
  60
     CUNTINUE
      STOP
      END
      SUBACUTINE LINV(P, N, T, FA, V, h)
      IMPLICIT MEAL * 8(A-H, K, L, O-Z), INTEGER (I,J,M,N)
      COMMON/INPUT/EPOCH(10), LAMBDA, EXPSVC, RHO, JTOL(8)
      COMMON/OUTPUT/QUAL(10), EXPWT(10,8), ITHTUT(8)
      LIMENSIGN E (25), & (25), H(25), V(50)
C. . ARGUMENT P = FUNCTION P (S) = LAPLACE TRANSFORM OF P(T)
             M = NC. OF VALUES OF S USED TO APPROAIMATE F (1)
             T = VALUE AT WHICH F (T) IS TO BE APPACXIMATED
             FA= AFPROXIMATE F(T) BASED ON N VALUES OF P(S)
             V = ABBAY OF COEFFICIENTS 'USED ADPEATEDLY)
             M = FORMAL FAMAMETER EXPLAINED BELOW.
```

C

C

C

```
ak IT & (6,44)
       FORMAT (1H , 351,2 (17HSTEHPEST'S METHOD, 201))
       WRITE (6,45)
      FORMAT (1H , 32X, 2 (25 ('-'), 18X))
       WEITE (6,46) N. N
      FORM AT (1H , 11X, 'T', 6X, 2('EXACT F(T)', 9X, 'M=', 12,9X,
      1'N=16',7X))
       WEITE (6,48)
  48 FORMAT (1H , 10 X, '----', 2(2X, 14('-')), 2X, y('-!),
      12(21,14('-')),21,9('-'))
       WEITE (0,50)
       FORMAT (1HO, 20X, "F (T) = 1/SQAS (PI*T) 1,24X,
      1^{\circ}F(T) = -C-Lh(T)^{\circ})
       WE IT & (6,52)
  52 FORMAT (1H , 16X, 2 (41 ('-'), 2X))
       DU 50 I=1.10
       WEITE (6,54) EPOCH (1), (EXACT (I,J), APPRCX (I,J),
      157 Enf (I,J), J=1,4,3)
      FORMAT (1H , 10X, F4. 1, 2(2(2X, F14.10), 2X, F9.5)).
      CONTINUE
       WRITE (6,60)
      10 \text{ kmat} (1 \text{ hO}, 26 \text{ X}, ^{\circ} \text{ F} (\text{T}) = (\text{T*T*T}) / 6^{\circ}, 27 \text{ X},
      1'F(T) = EXF(-T)'
       WRITE (6,52)
       ₩0 66 I=1,10
       White (6,54) LPOCH (I), (EXACT (I,J), APPROX (I,J),
      151 EHF (1,J), J=2,5,3)
      CONTINUE
       WE ITE (0,70)
  70 FORMAT (1HO, 26X, F(T) = SIN (SURT (2+T)) 1,22%,
      1'F(T) = 1-3+T+3+T+T/2-T+T+T/6'
       WEITE (0,52)
       DO 70 I=1,10
       WAITE (6,54) EPOCH (I), (EXACT (I,J), APPROX (I,J),
      1STEHF (1,J), J=3,6,3)
  76
       CUNT INUL
      CUNTINUE
  04
       STOP
       EN D
       SUBACUTINE LINV(P, N, T, FA, V, h)
       IMPLICIT REAL+8(A-H, K, L, O-Z), INTEGER (I,J, H, N)
       COMMON/INPUT/EPOCH(10), LAMBDA, EXPSVC, RHO, JTOL(8)
       CUMMON/OUTPUT/QUAL(10), EXPWT(10,8), ITHTUT(8)
       LIMENSIGN E (25), & (25), H(25), V(50)
C. ARGUMENT P = FUNCTION P(S) = LAPLACE TRANSFORM OF P(T)
              h = NC. OF VALUES OF S USED TO APPROALHATE F (1)
              T = VALUE AT WHICH F (T) IS TO BE APPAGNIMATED
              FA= APPROXIMATE F(T) BASED ON N VALUES OF P(S)
              V = ABBAY OF COEFFICIENTS 'USED ABPEATEDLY)
              M = FORMAL FAMAMETER EXPLAINED BELOW.
```

C

```
II (M. L. M) GO TO 50
C. ON FIRST CALL OF LINV (M NOT EQUAL TO N) AMMAY V(I) HUST
 DE EVALUATED.
C. FIRST COMPUTE E(I) = FACTORIAL OF 21 OVER INOSE OF I AND
 I-1, AND B(I) = FACTORIAL OF I-1.
      MH = N/2
      £(1) = 2
      E(2) = 12
      F(1) = 1
      F(2) = 1
      DG 16 1 = 2, NH
      E(I+1) = (2.D0*(2*I+1)*E(I)) / I
      F(I+1) = F(I) + I
  10 CONTINUE
C. . BEXT COMPURE H (J) = ALL TERMS INDEPENDENT OF INDEX I IN
C TEE V(I) SUMMATION.
      PWA = NH
      H(1) = 2/F(NH)
      DO 20 J = 2, MH
      H(J) = \{(J + + y + R) + E(J)\}/Y(H + J + 1)
  20 CONTINUE
C. . FINALLY EVALUATE ARRAY V(I) .
      ISIGN = 2*(NH-2*(NH/2))-1
C
      10 40 I = 1, M
      V(I) = 0
      JFIRST = (I+1)/2
      JLAST = MINO(I,NH)
C
      DO 30 a = JEIRST, JLAST
      V(I) = V(I) + H(J)/(F(I-J+1)*F(2*J-I+1))
  30
      CONTINUE
      V(I) = ISIGN*V(I)
      ISIGN = -ISIGN
  40
     CONTINUE
      M = b
C. . SUBSEQUENT CALLS (K EQUAL TO N) WILL USE ARRAY V(I) PROM
  COMMON STORAGE AND WILL BRANCH TO THIS SECTION OF LINV.
  50 FA = 0
      A = .693147180559945309417232121458D0 / 1
```

```
CO 60 I= 1, h
FA = FA + V(I) *P(I*A)
  60
      CONTINUE
      FA = A*BA
C
      BETURN
       END
C
C
      DOUBLE PRECISION FUNCTION P1 (S)
      IMPLICIT REAL+8 (A-H, O-Z), INTEGER (I-N)
       £1=1/DS_AT(S)
      RETULL
       END
      DOUBLE PRECISION FUNCTION P2(S)
      IMPLICIT MEAL+8(A-4, 0-2), IMTEGER (I-N)
       £2=1/(5*5*5*5)
      EETUAN
      END
C
C
      LOUBLE PRECISION FUNCTION P3(5)
      IMPLICAT REAL+8(A-H, O-Z), INTEGER (I-N)
      P3=DSQ#T (3. 1415926535897932384626433/(2*S*S*S))
      1+UEX & (-1/(2+5))
      AETULA
      END
C
      DOUBLE FRECISION FUNCTION F4(S)
      IMPLICIT LEAL+ 6(A-H, G-Z), INTEGER (I-N)
      £4 = DLUG (S) /S
      BETUEN
      END
C
C
      DOUBLE PRECISION FUNCTION PS(5)
      IMPLICIT REAL+8(A-H, U-Z), INTEGER (I-N)
      £5=1/(S+1)
      BETUAN
      END
      DOUBLE PRECISION FUNCTION PO(S)
      IMPLICIT BEAL+8(A-H,O-Z), INLEGEA (I-N)
      P6=(S-1)*(S-1)*(S-1)/(S*S*S*S)
      KETUEN
      ED C
```

```
COMPUTER PROGRAM FOR FIGURE 3.4
      IMPLICIT MEAL+8(A-m,0-Z), INTEGER (I-N)
      COMMCA PA1 (10) , PA2 (9)
      LIMENSIGN F1 (10), 51 (10), D1 (10), V1 (10),
     112 (9) , 62 (9) , V2 (9) , V (50) ,T2 (9)
      EXTERNAL #1.F2
      M=0
      N=34
      DC 2C 1=1,10
      £1 (1) = 1/DSQRT (2+3.1415926535697932364620433)
      CALL LINV(P1, N,T, FA1(I),V, M)
  20
      CONTINUE
      DATA 51/.50555,.39912,.32655,.28276,
     1.25174,.22989,.21322,.19956,.18814,.17790/
      LATA U1/.73172,.40035,.26343,.26286,
     1.29305,.22901,.18062,.20112,.21069,.17650/
      DATA V1/.56419,.39894,.32573,.28209,
     1.25231,.23033,.21324,.19947,.18666,.17841/
      DATA T2/4.140186, 2.501126, 1.643438, 1.085064,
     1.693147,.412298,.214821,.085541,.016048/
      DC 30 1=1,9
      1=12 (I)
      F2 (1) = DEXP (-1/2)
      CALL LINV (P2, N,T, PA2 (I), V, M)
  30
      CONTINUE
      DATA 62/. 120527, . 288195, . 439084, . 581306,
     1.707316,.813401,.898482,.957847,.992205/
      DATA V2/.126174,. 286329,. 439675,.581265,
     1.707107,.813712,.898156,.958135,.992015/
C
      DO 180 MCOPY=1,5
      akITE (0 .4)
   4 FOLHAT (1H1, ////)
      MARGIN = 4 + MCOPY
      DO 8 IMAGN = 5, MARGIN
      #RITE (0,0)
     FORMAT (1H )
     CONTINUE
      WRI2E (6,40)
      FORMAT (1H ,33X, FIGURE 3.4 COMPARISONS',
     1. WITH CTHEL TECHNIQUES')
      WR 11 1 (6,48)
  46 FORM AT (1H , 33 X ,45 ('-'))
      #RITE (6,52)
  52
      FORMAT (1HO, 20X, "F (T) = 1/SQRT (T*FI) ")
```

```
WRITE (0,54)
     FORMAD (1H , 20X, 19 ('-') , 3X, 'APPROXIMATE F(1) !.
    1' USING NUMERICAL INVERSION')
     BEITE (0,60)
     FORMAT (1H , 42X, 42 ('-'))
     WRITE (6,64)
     FORMAT (1H , 44 X, 17 HSTEHFEST'S METHOD)
     white (6,66)
     FORMAT (18 , 42x, 21 ('-'))
     WEITE (6 .70) N
     FORSAT (14 , 22X, T
                              EXACT F (T)
                                                 M=1,12,6X,
    1'h=10*
                  DUB NEA *
                              VEILLON* )
     #BITE (6,72)
 72 FORHAT (1H , 21 X, '---', 2 (4X, 10 (!-')), 3 (4X, 7 ('-')))
     DU 78 I=1,10
     WRITE (6,76) 1,F1(I),FA1(I),S1(I),D1(I),V1(I)
     FORM AT (1H , 21X, 12,5X, 2 (F10.8, 4X), 3 (F7.5,4X))
 76
 78
     CONTINUE
     (08, 3) 4TLEN
     FORMAT (1h-, 20X, P (T) = EXP (-1/2) )
 BC
     wRITL (0_82)
     FORMAT (1H , 20X, 16 ('-'), 10X, 'APPROXIMATE F(T) ',
    1' USING MUMERICAL INVERSION')
     WE IT E (0,85)
    TORNAT (1H , 46X , 42 ( - 1) )
     WAITE (6 ,86) N
     FORMAT (1d , 24X, T', 9X, 'EXACT F (T) ', 6X,
    1'STEHFEST M=', I2, 4X, BELLHAM*
                                         VEILLON* ')
     wRITE (0,80)
 88 FORMAT (1H , 20x,8('-'),2(3x,15('-')),2(3x,8('-')))
     DU 94 I=1,9
     WE ILE (6,92) 12 (1), F2 (1), FA2 (1), B2 (1), V2 (1)
     FORM AT (1H , 20 X , 28.0, 2 (3X , F 15. 13) , 2 (3X , F 8.6) )
 92
     CONTINUE
 54
     WEITE (6, 102)
     FURNAT (1H-, 25%, * MOTE: DATA FUR BETHODS MARKED .
    1' BY AN ASTELISK WERE TAKEN')
     WE ITE (6, 104)
    FORMAT(14 , 32X, FROM A.C. M ALGORITHM 400, ",
    123H MUMERICAL INVERSION OF)
     wk ITE (0, 106)
     FORMAT (1H , 32X, 23HLAPLACE THANSFORM', BY ,
    1'FRANCGISE VEILLUN, COMMUNI-')
     WRITE (6, 108)
     FORMAT (1H , 32X, CATIONS OF THE A.C.M., ,
    1' VOLUME 17, NUMBER 10,")
     WRITE (6, 110)
110
     FURNAT (1H , 32X, 'UCTOBER 1974.')
100
     COMILAUE
     STUP
     ENL
```

```
SUERCULINE LINV(P,N,T,FA,V,d)
    IRPLICIT MEAL #6(A-H, O-Z), INTEGER (1-M)
    LOMBUL FA1 (10) , FA2 (9)
    DIMENSION E(25), F(25), H(25), V(50)
   IF (A.E.L.) GO TO 50
    MH = M/2
    £(1) = 2
    E(2) = 12
    £(1) = 1
    £ (2) = 1
    10 10 1 = 2, MH
    E(I+1) = (2.00*(2*I+1)*E(I)) / I
   F(I+1) = F(I) *I
10
   CONTINUE
    FWH = MH
    H(1) = 2/E(NH)
    DO 20 J = 2, NH
    H(J) = ((J**PHE)*E(J))/F(HH-J+1)
20 CONTINUE
    ISIG# = 2*(NH-2*(NH/2))-1
    DO 40 I = 1, N
    V(I) = 0
    JFIRST = (1+1)/2
    JLAST = MINO(I,NH)
    DO 30 J = JFIRST, JLAST
    V(I) = V(I) + H(J)/(F(I-J+1)*F(2*J-I+1))
30
   CONTINUE
    V(I) = ISIGN + V(I)
    ISIGN = -ISIGN
40 CONTINUE
    M = N
50
    A = .69314718055 y 94530 94 17232121458DQ / T
    DO 60 I= 1, N
    FA = FA + V(I) *P(I*A)
   CONTINUE
    FA = AFEA
    KETUAN
    END
    DOUBLE PRECISION PUNCTION P1(S)
    IMPLICIT MEAL+8 (A-b, 0-Z), IMPLGED (I-M)
    £1=1/D5QR1(5)
    RETURN
    EN D
    DOUBLE PRECISION FUNCTION P2(S)
    IMPLICIT MEAL+8(A-H, O-Z), INTEGER (I-M)
    P2=4/(1+2+5)
    SETUEN
    END
```

COMFUTER PAGGRAM FOR FIGURE 3.5

```
IMPLICIT REAL+8(A-H, K, L, O-Z), INTEGER (I,J, M, N)
      COMMCH/INFUT/EPOCH(10), LAMBDA, EXPSVC, RHU, JTUL(8)
      COMMCH/QUTPUT/QUAL(10), EXPHT(10,8), ITHTOT(8)
      CUMBUL/LAPLAC/V(50), M, N
      CUMMER/ME 1/OSTART,SS
      COMMON/MEWRAP/TOLANC, ITHSUM
      EXTERNAL FUAD, PAM1
      EXTERNAL FAMI, DF MMI, DUPMMI
      t. = 0
      h = 34
              = 0.00
C
      DC 20 I = 1,10
      CALL LINV (PQUAD, N , EPOCH (I) , EZIT, V, M)
      QUAD(I) = (BHC-1.) *EPOCH(I) + EZIT
      CONTINUE
  20
      DO 50 J=1,8
      JICL (J) = 2*J + 8
      TOLENC = 1./(10.**JTOL(J))
      OSTARL = 1
      IT MSUM = C
      DO 46 1=1,10
      CALL LINV (PMM 1, N, EFOCH (I) , EZIT, V, M)
      EXPWT(I,J) = (EHO-1.) * EPOCH(I) * EZIT
  40
      CONTINUE
      ITATUT(J) = ITASUM
C
      CUNT INUL
  50
C
      CALL TOLUUT
      STOP
      END
      BLOCK DATA
      INFLICIT REAL *8 (A-H, K, L, G-Z), INTEGER (I, J, M, M)
      COMMON/IMPUT/EPOCH (10) ,LANBOA , EXPSVC, EHC, JTOL (8)
      DATA EFGCh /20., 40., 60., 80., 100.,
     1200.,300.,400.,800.,1200./
      DATA LAMBDA, ELPS VC, RHO /1.00,0.95,0.95/
      END
      SUBROUTIAL TOLOUT
      IMPLICIT MEAL+8(A-H,K,L,O-Z), IMTEGER (I,J,H,N)
      COMMON/INPUT/EFOCH (10) , LAMBLA, EXPSVC, EHO, JTOL (8)
      COMMUN/CUTPU1/QUAL(10), EXPNT (10,8), ITNTU1(8)
```

```
CUMMUN/LAFLAC/V (5G), M, M
      LIMENSIUM BESSEL (5) , JTOL (8)
      BESSEL (1) = 3. 9161
      EESSEL (2) = 5.4651
     BESSEL(3) = 6.5582
      BESSEL (4) = 7.4191
      BESSEL (5) = 8. 1335
      DU BL NCOPY=1,5
      WAITE (6,4)
    FORMAT (1H1,////)
      MARGIN = 4 + NCOPY
      DO 8 INEGN = 5,MARGIM
      #RITE (6,6)
  b FORBAT (1H )
   8 CONTINUE
      #KITE (6,14)
   12 FORMAT (1H , 29X, 'PIGURE 3.5 EFFECT OF NEWTON-RAPHS CK',
     1' SEARCH TOLERANCE')
      MAITE (6,14)
   14 FORMAT (1H , 29X,53 ('-'))
      #RITE (0, 10)
   16 FORMAT (1HO, 1CY, 37 HIHIS TABLE COMPARES COLEMAN'S BESSEL .
     1' FUNCTION RESULTS AT FIVE EPOCHS WITH LAPLACE')
      WRITE (6, 18)
   16 FORMAT (1H , 10X, 'INVERSION RESULTS FOR BOTH THE EXACT ',
     1 QUADRATIC SOLUTION OF THE FUNCTIONAL AND THE ')
      #RITE (6, 20)
   20 FORMAT (1H , 10X, APPROXIMATE NEWTON-RAPHSON SEARCH ',
     1. SOLUTION OF THE FUNCTIONAL WITH VARIOUS TOLERANCES. 1)
      WRITE (6,30)
   30 FURHAT (1m0, 10x, d/M/1 MEAN SERVER LOAD, ..
     1' LAMBDA = 1.0, E(S) = 0.95, 4 = 0.1)
C
      DO 8C JEIKS1=1,5,4
      JLAST=JFIRST+3
C
      wRITE (6,36) M
   38 FORMAT (1HO, 36X, 'INVERSION OF LAPLACE THANSFORM, ',
     122H STEHFLST'S METHOD, N=,12)
C
      MAITE (0,40)
   40 FURNAT(1m , 18X, 'SUES OF ',73('-'))
      WE ITE (6,42)
   42 FORMAT (1H , 19X, BESSEL QUADRATIC
                                                TOLERANCE FOR .
     1 'NEWTON-MAPHSON SEARCH SOLUTION OF PURCTIONAL')
```

```
C
      "BITE (6,44)
   44 FORMAT (18 , 10x, EFOCH PUNCTIONS SOLUTION OF
                                                            1,58(1-1)
C
      DO 46 J=JFIRST,JLAST
  46 CONTINUE
C
       MEITE (6,50) (JTOL (J), J=JFIEST, JLAST)
   50 FORMAT (1H , 10X, T
                                (COLLMAN)
                                             FUNCTIONAL
     14('10++-', I2, 8X))
C
       WRITE (0,56)
   56 FORMAT (1H , 10x, '----
                                  -----',5(2x,13('.-').))
C
      DO 62 I=1.5
C
      WAITE (0,60) EPOCH (I) , BESSEL (I) , UAD (I) ,
     1 (EXPWT (I, J), J=JFIBST, JLAST)
      FORMAL (1H , 10X,F5.0, 3X,F7.4,5(1X,F14.11))
  t2
     CUNTINUE
      LO 60 1=6,10
C
      WRITE (6,64) EPOCH (I), QUAD (I), (EXPHT (I, J), J=JPIRST, JLAST)
      FORMAT (1H , 10X, F5.0, 10X, 5 (1A, F14.11))
  66
      CONTINUE
C
      WEITE (6.70)
   70 FORMAT (1H , 10X, 'TOTAL NUMBER OF NEWTON-RAPHSON')
C
      *AITE (6,72) (ITHTOT (J), J=JFIEST, JLAST)
   72 FOREAT (1H , 10X, ITERATIONS TO EVALUATE 10 EPOCHS:
     14 (I4, 11X))
   80 CUNTINUE
      KETUKN
      ENL
C
      SUBROUTINE LINV (P. N. T. FA, V. A)
      IMPLICIT REAL+8 (A-H, U-Z), INTEGER (I-N)
      COMMEN/LAFL AC/EPO CH (10) , APPROX (10, 6).
      DIMENSION E (25), E (25), H(25), V (50).
      IF (M. Eu. M) GO TO 50
      MH = M/2
      E(1) = 2
      E(2) = 12
      F(1) = 1
      F(2) = 1
      UU 10 I = 2, NH
      z(1+1) = (2.00+(2+1+1)+E(1)) / 1
      E(1+1) = F(1)+1
```

```
10
      CONT INUE
      FWA = AH
      11(1) = 2/E(NH)
      DC 20 J = 4, MH
      H(J) = \{(J**2WE)*E(J)\}/F(WH-J+1)
  20
      CONTINUE
      ISIGA = 2*(NH-2*(NH/2))-1
      DO 40 I = 1, N
      V(I) = 0
      JFIRST = (1+1)/2
      JLAST = MINO(I,NH)
      DO 30 J = JEIEST, JLAST
      V(I) = V(I) + H(J)/(F(I-J+1)*P(2*J-I+1))
  30
      CUNTINUE
      V(I) = ISIGN*V(I)
      ISIGN = -ISIGN
  40
      CONTINUE
      M = N
      FA = C
  50
      A = .693147180559945309417232121458DC / T
      DO 60 I= 1, N
      FA = FA + V(I) *P(I*A)
      CONTINUE
      FA = A*EA
      ALTUKA
      END
C
      MUBLE PRECISION FUNCTION PUND (S)
      IMPLICIT REAL+8 (A-b, K, I, O-Z), INTEGER (I, J, M, M)
      COMMUN/INPUT/EPOCH (10) ,LAMBUA, EXPSVC, BHC, JTOL (8)
      B = 1 - EXPSVC* (LAMBDA+S)
      LMEGA = (-b+LSQET(b+B+4+S+EXPSVC))/(2+EAPSVC)
      LUUAD=DEXP (-OMEGA+Z) / (S+OMEGA)
      RETURN
      END
      DOUBLE PRECISION FUNCTION PANT (S)
      IMPLICIT REAL #8 (A-u, K, L, U-Z) , INTEGER (I, U, A, M)
      CUMMUN/INFUT/EPOCH (10), LAMBDA, EXES VC, RHO, JTO1(8)
      COMMON/MAT/OSTART,SS
      EXTERBAL FCMa1, DFMM1
      55 = 5
      CALL ZERC (FOMM 1, DEMM 1, OSTART, OMEGA)
      OSTABL = OMEGA
      PMM1=DEXP(-CMEGA+Z)/(S+OMEGA)
      RETURN
      END
      LOUBLE PARCISION FUNCTION FAM1 (S)
      IMPLICIT MEAL+8 (A-H, K, L, O-Z), INTEGER (I, J, A, h)
      COMMON/IMPU1/EPOCH (10) ,LAMBDA , EXPSVC, RHC, JTOL (8)
```

```
FMm1 = S - LAMBDA* (1- (1/(1+5*EXPSVC).))
      bETU ba
      END
C
C
      DOUBLE PRECISION FUNCTION DPMM1(5)
      IMPLICIT REAL+8(A-H, K, L, 0-Z), INTEGER (I,J, M, M)
      COMMON/INPUT/EPOCH (10) , LAMBOA, EXPSVC, RHO, JTOL (8)
      DFMM1 = 1 - LAMBDA*EXPSVC* (1/ (1+5*EXESVC)) **2
      RETURN
      END
      DOUBLE FRECISION FUNCTION DDFMM1 (S)
      IMPLICIT MEAL+ 6(A-H, K, L, O-Z), INTEGER (I,J,M,N)
      COMMON/INPUT/EPOCH(10), LAMBDA, EXPSVC, RHO, JTOL(8)
      DDFMM1 = 2*LAM EDA*EXPSVC**2*(1/(1+5*EXPSVC)) **3
      BETUAN
      EAD
      DOUBLE PRECISION PUNCTION FORM 1 (CHEGA)
      IMPLICIT AEAL+8 (A-H, K, L, O-4), INTEGER (I, J, a, h)
      COMMON/Ma1/OSTAKT,SS
      FUME1 = FMM L(ONEGA) - SS
      BETURN
      END
C
      SUBBOUTINE ZERO (F, DF, X START, AZERO)
      INFLICIT REAL+8 (A-h, K, L, O-Z), INTEGER(I,J,a, N)
      COMMON/MEWBAR/TOLHNC, ITHSUM
      11 ERAT = 0
      X = XSTART
  10
      CONTINUE
      XNEW = X - F(X)/\nu F(X)
      MELACC = (XNEW-X) /XNEW
      A = ANEW
      ITERAT = 11 ERAT + 1
      IF (ITERAT .GE. 50) GO TO 50
      IF (LABS (RELACC) .LE. TOLRNC) GO TO 20
      GU TU 10
      CUNTINUE
      ITASUK = ITASUM + ITERAT
      AZERC = X
      BETULN
  50
      CONTINUE
      MEITE (6,55) RELACC, TOLANC
      FORMAT (1H , 'ITERAT=50, RELACC=', D24.16,', TOLENC=', D24.16)
      ITASUM = ITASUM + ITERAT
      XZERO = X
      BETURN
      EN D
```

```
COMPUTER PROGRAM FOR FIGURE 3.6
      IMPLICIT MEAL+8 (A-E, U-Z), IMTEGER (I-N)
      CUENCH /AHOZ/ Bud, 2
      DIMENSION SEZIT1 (11,8), SEZIT2 (11,8), V (50), ZSIN (8), MICH (11)
      EXTERNAL PUULU1, PQUAD2
      DATA ZSIA /.2,.4,.6,.8,1.,2.,3.,4./
      DATA BEIGIN /.5,.0,.7,.8,.9,1.1,1.2,1.3,1.4,1.5,2./
C
      M=0
      N=34
C
      DO 12 IBHO=1,11
      RHO = AHOIN (INHO)
      WIFCIA = LABS (1. -RHO)/RHO
      DO 12 JZ=1,8
       Z = ZSIN(JZ)/WTFCTE
      CALL LINV (PQUAD1, N, Z, EZIZ, V, A)
       SEZIT1(IRHO,JZ) = EZIZ##TFCT#
      CALL LINV (PUNCZ, N.Z, EZIZ, V.M)
      SEZIT 2 (IMMC, JZ) = EZIZ # WTFCTA
      CONTINUE
  12
       DO 88 NCUPY =1.5
       WRITE (6, 20)
      FORM AT (1H1,////)
       MARGIN = 4 + NCOPY
       DO 28 IMEGN = 5, MARGIN
       WE ITE (6, 24)
      FORMAT (1H )
  28
      CONTINUE
       WE ITE (6, 32)
      FURNAT (1H , 26X, FIGURE 3.6 CORRECTIONS',
      1' FOR DISCONTINUOUS FIRST DERIVATIVE')
       #RITE (6, 36)
      FORMAT (1H , 20X, 50 ('-'))
       WHITE (6,40) N
      FORMAT (1H-, 33X, APPHOXIMATIONS USING',
      122H STEHFEST'S METHOD, N=,12)
       WRITE (6,44)
      FORMAT (1H-, 20%, " M/M/1 SCALED MEAN CUMULATIVE",
      141H IDLEMESS AT T'=ABS (MU) *2' (I.E., AT T=2))
       WRITE (6,48)
      FORMAT (1HO, 47X, WITHOUT CORRECTION TERM')
       WRITE (0,52)
  52
       FORMAT (1HU, 45%, SCALED INITIAL SERVER LUAD')
       WAITE (6,56)
       FORMAT (1H , 19x,76 ('-'))
       WE ITE (6,00) (ZSIM (JZ), JZ=1,6)
  60 FURNAT (1H , 15X, 's HU', 5X, 8 (F3. 1, 7X))
       #RITE (6,64)
       FORMAT (1H , 15x, '---', 8 (2x, 8 ('-')))
  64
```

```
DO 72 IRHO=1,11
      WEITE (6,68) RHOIN (INHO), (SEZIT1 (1MHO, J4), J4=1,8)
  68
      FORM A1 (1H , 15%, #3. 1,8 (2%, P8.0))
  72
      CONTINUE
C
      WA ITE (6,76)
  76
      FORMAT (1H )
      WRITE (6,44)
       WE IT'E (6,80)
      FORM AT (1HG, 48%, WITH CORRECTION 1ERM')
      WEITE (6,52)
      WEITE (6, 56)
      WEITE (0,00) (ZSIN (JZ), JZ=1,8)
      WEITE (6,64)
      DO 84 IEHO=1,11
      MAITE (6,68) AHOIN (THO), (SE4112 (IRHO, JZ), JZ=1,8)
  84
      CONTINUE
      CONTINUE
      STOP
      EN D
C
C
      SUBACUTINE LINV (P. N. T. FA, V.A)
      IMPLICIT BLAL* 8 (A-H, O-Z), INTEGER (I-N)
      COMMON/LAPLAC/ EPU CE (10) , APPROX (10, 6)
      DI demoion & (25), F(25), H(25), V (50)
      IF (M.Eu. M) GO TO 50
      MH = 6/2
      L(1) = 2
      E(2) = 12
      F(1) = 1
      £ (2) = 1
      DO 10 1 = 2, NH
      E(I+1) = (2.00*(2*I+1)*E(I)) / I
      £(1+1) = £(1)+1
  10 CONTINUE
      Pan = NH
      H(1) = 2/F(NH)
      DC 20 J = 2, NH
      H(J) = ((J++PHE) + E(J)) / E(H-J+1)
  20
      CONTINUE
      ISIGN = 2*(NH-2*(NH/2))-1
      DO 40 1 = 1, N
      V(I) = 0
      JF IBS1 = (I+1) /2
      JLAST = MINO(I, Mh)
      DO 30 J = JEIEST, JLAST
      V(I) = V(I) + H(J)/(F(I-J+1)*F(2*J-I+1))
      CONTINUE
  30
      V(I) = ISIGN+V(I)
      ISIGN = -ISIGN
  40 CUNTIBUE
      H = #
```

```
50 FA = 0
    A = .69314718055994530941723212145800 / T
    DO 60 I= 1, N
    EA = EA + V(1) *P(1*A)
   CONTINUE
    FA = A+FA
    RETURN
    END
    DOUBLE PRECISION FUNCTION FRUADI (S)
    IMPLICIT MEAL+8 (A-b, 0-2), INTEGER (I-M)
    COMMUN /RHOZ/ RHO,2
    B = 1 - BHO - 5/2.
    OMEGA = -8+D5QRT (B*B+2*S)
    PUUAL1 = DEXF (-OMEGA+Z)/ (S+UMEGA)
    PETO BP
    END
    DOUBLE PRECISION FUNCTION PUUND2 (S)
    IMPLICIT MEAL *8 (A-H, O-Z), INTEGER (I-N)
    COMMON /RHOZ/ RHO, Z
    B = 1 - RHO - 5/2.
    CALGA = -B+ DSQRT (B+B+2+S)
    PQUAD2=DEXF(-OMEGA*Z)/(S*S)+DEXP(-Z*(2.*AHO+S))/(S*S)
    AETUAN
    END
```

```
COMPUTER PROGRAM FOR APPENDIX TABLES, ENO NOT EQUAL TO ONE
       IMPLICIT REAL *8 (A-H, K, L, O-Z), INTEGER (I, J, H, N)
      COMMUN /INGUT/ TVAL (16) , KVAL (7) , SCL HT (10 , 16 , 3) , EHOIN,
      125IN (3)
      COMMON /LAPLAC/ V (50) , A, N
      COMMON /NEWRAP/ TOLENC, MAXITM
C
      M=C
      w= 34
      MAXITH= 20
      10LANC= 1. D- 20
C
      DO 60 I EAGE = 1, 3
      READ (5, 20) RHOIN, (2SIN (ITABLE), ITABLE=1,3)
      KEAD (5, 20) (TVAL (3), J=1,6)
       MEAD (5, 20) (TVAL (J), J=7, 12)
      EEAD (5, 20) (TVAL (J), J=13, 16)
     PUMAT (6 F10.0)
  20
      DC 50 ITABLE=1.3
      CALL EZWNE (ITABLE)
      CALL MAMD1 (ITABLE)
      CALL AZMEK (ITABLE)
      CALL RZGAM (ITABLE)
      CONTINUE
      LALL AZOUT
  60 CONTINUE
      STOP
      EN D
      IMPLICIT REAL*8 (A-H, K, L, O-Z), INTEGER (I,J, M, N)
      COMMCH /INCUT/ TVAL(16), KVAL (7), SCLHI (10,16,3), LHO IN,
     125 IN (3)
      CATA KVAL/4., 3., 2., 1.5, 1., .5, .2/
      ENL
C
      SUBROUTINE LINV (P. h. T. FA, V. M)
      IMPLICIT REAL+8 (A-H, O-Z), INTEGER (I-N)
      COMMON/LAPLAC/EPOCH(10), APPAGE (10,6)
      DIMERSION E (25), F (25), H(25), V (50)
      IF (a. Ey. b) GO TO 50
      NH = M/2
      E(1) = 2
      E(2) = 12
      £(1) = 1
      F(2) = 1
      LO 16 1 = 2, MH
      E(I+1) = ( 2.00*(2*I+1)*E(I) ) / I
      E(1+1) = E(1)+1
```

```
10
    CONTINUE
    PHE = AH
    H(1) = 2/F(MH)
    DO 20 0 = 2, MH
    H(J) = ((J**PHE)*E(J))/F(HH-J+1)
20
   CONTINUE
    ISIGN = 2*(NH-2*(NH/2))-1
    DO 40 I = 1, h
    V(I) = 0
    JFIR52 = (1+1)/2
    JLASA = MING(I,NH)
    DO 30 J = JFIAST, JLAST
    V(I) = V(I) + H(J)/(F(I-J+1) *P(2*J-I+1))
   CONTINUE
30
    V(I) = ISIGN*V(I)
    ISIGN = -ISIGN
40
   CONTINUE
    a = M
50
   FA = 0
    A = .693147180559945309417232121458D0 / 1
    DO 60 I= 1, M
    FA = FA + V(I) *P(I*A)
   CONTINUL
    FA = A*FA
    LETUEN
    END
    SUBRULTIME MAUNA (ITABLE)
    IMPLICIT REAL+8 (A-b, K, L, O-2), INTEGER (I,J, M, h)
    COMMEN /INCUT/ TVAL(16) ,KVAL(7) ,SCLWT (10, 10, 3) , RHOIM,
   125 In (3)
    COMMON /LAFLAC/ V (50), M, N
    COMMCN /AHOZO/ AHO, 2, OSTART
    EXTERNAL FUND
    SCALING WITH SIGSON = 1
    BHC = BHCIN
    ALFS(R = (1-EHC) + (1-EHO)
    WIFCIR = DABS (1-AHO)
           = 25 IN (ITABLE) / WIFCTR
    DO 20 J=1,16
    I = TVAL (J) /ALESUE
    CALL LINV (PUNR, N, T, TERMS, V, M)
    EZWT = Z + (AHO-1.) *T + TERM3
    SCINT (1, J, ITABLE) = EZ HT* WT FCTA
    CONTINUE
    BETUHN
    LND
```

```
C
      DOUBLE PRECISION FUNCTION PHNE (S)
       IMPLICIT REAL+8(A-H, K, L, O-Z), INTEGER (I, J, H, N)
      CUMMON /AHU 20/ AHC, Z, OSTART
      Chega = -(1-kho) + DSQRT ((1-kho) + (1-kho) + 2*s)
      PWNR = DEXP (-OMEG A+Z) / (S+OMEGA)
      BETUAN
       ENU
      SUBLOUTINE REMLT (ITABLE)
       IMPLICIT MEAL+8 (A-H, K, L, O-2), INTEGER (I, J, H, H)
      COMMON /INCUT/ TVAL (16) , KVAL (7) , SCLHT (10, 16, 3) , RHOIN,
      125 In (3)
       CUMBON /LAPLAC/ V (50), M, N
      COMMON /MHOZO/ RHO, Z,OSTART
       EXTERNAL PMD1, FMD1, DPH C1, DUEND1
C
C
      SCALING WITH SIGSON = RHO
      RHC = BHOIN
       ALFS _H = (1-RHO) * (1-RHO) /RHO
       WTFCTR = DABS (1-RHO) / HHO
       OSTART = 1
              = 25IN (ITAELE) / WTFCTE
       ELZ = DEXP (-KHO+Z)
       IF (BHC .LE. 1.C) GO TO 10
       CALL ZEHO (DPMD1, DDFM D1, 10-10, SBAK)
       DELTA = -PMD1 (SBAB)
       START = SBAR + DELTA
       CALL LENG(F MD1, DFMD1, START, OSTART)
  10
      CONTINUE
       DO 20 J=1,16
       T = TVAL(J) /ALFSQR
       IF (1 . LE. Z) EZI1 = 0.
       IF (T .GT. 2) CALL LIAV (PADI, M.T. HT, V.M)
       IF (1 .61. 2) EZIT = HT + (I-Z) + ELZ
       EZ # 2 + (RHO-1.) * T + EZIT
       SCL WT (2,J, ITABLE) = EZWT*WTFCTA
       CONTINUE
       BETURN
       END
C
```

```
LOUDLE PARCISION FUNCTION PAD1 (S)
      IMPLICIT REAL+6 (A-H, K, 1, 0-2), INTEGER (I,J, M, h)
      COMMEN /AHUZU/ RHU, Z, OSTART
      COMMEN SS
      EXTERNAL FORL 1, DE MD1
      SS= S
      CALL ZEAD (FOMD1, DEMD1, OSTART, OMEGA)
      OSTABY = CHEGA
      Ph D1 = DEXP (-OMEGA=2)/(S+OMEGA) - DEXP (-2+ (RHC+S))/(S+S)
      BETULN
      END
      DOUBLE PRECISION FUNCTION PAUL (S)
      IMPLICIT MEAL +8 (A-H,K,L,O-Z), INTEGER (I,J,A,N)
      COMMCA /RHCZO/ EHO, Z, OSTART
      FMD1 = S-RHO* (1-DEXP (-S))
      BETULL
      LAD
C
      DOUBLE PRECISION FUNCTION DENDI(S)
      IMPLICIT REAL+8(A-H,K,L,O-Z), INTEGER(I,J,M,K)
      COMMON / MHOZO/ MHO, Z, OSTART
      LFMD 1 = 1-RHO+DEXP (-S)
      BETURL
      LND
C
      DOUBLE PRECISION FUNCTION DUFAD1 (S)
      IMPLICIT REAL+8(A-H,K,L,O-Z), IMTEGER(1,u,K,N)
      COMMON /RHG23/ RHG, 2, OSTART
      LDFAU1 = AHO*DEXP (-S)
      BETUEN
      END
      DOUBLE PRECISION FUNCTION FOAD1 (OMEGA)
      INFLICIT ME AL +8 (A-H,K,L,O-Z), INTEGER (I,J, H, A)
      COMMON SS
      FOND1 = caul(OMEGA) - SS
      BETUEN
      END
```

```
C
      SUBMOUTINE BZMEK (ITABLE)
      IMPLICIT REAL+8(A-H, K, L, O-Z), INTEGER (I, J, M, N)
      COMMON /INGUT/ TVAL(16), KVAL(7), SCLHT(10,16,3), RHOIM,
      1251# (3)
      COMAGA /LAPLAC/ V (50) . M. N
      COMMEN / HHO 20K/ HHO, Z, OSTABT, K
      LITEBRAL PMEK, FMEK, DPMEK, DDFMEK
      SCALING WITH SIGSON = (K+1) *AHO/K
      PHC = FHOIR
      DU 30 1=3, 9
               = KVAL (I-2)
      ALFSQH = K* (1-RHO)* (1-RHO)/((K+1)*AHO)
      wTFCTk = k*DABS(1-kHO)/((k+1)*RHC)
      USTART = 1
             = 2SIm (ITABLE) / STFCT&
      ELZ = DEAP (-RHO+ 2)
      IF (RHO .LE. 1.0) GO TO 10
      CALL ZENO (DIMEK, DDFMEK, O, SBAM)
      LELYA = -FMEK (SBAR)
      START = SBAR + DELTA
      CALL ZERO (FMEK, Dr MEK, START, OSTART)
  10
      CONTINUE
      DO 20 J=1,16
      1 = TVAL (J) /ALFSUR
      IF (1 .LL. 2) EZIT = 0.
      IF (T .GT. Z) CALL LINV (PMEK, M, T, HT, V, H)
      IF (1 .GT. 2) EZIT = HT + (T-2) * ELZ
      EZWT = 2 + (aHC-1.) *T + EZIT
      SCLWT (I, J, ITABLE) = EZWT*WIFCTR
  20
      CONTINUE
  30
      CUNTINUE
      AETUHN
      END
      DOUBLE PRECISION FUNCTION PHEK (5)
      IMPLICIT REAL *8(A-H,K,L,O-Z), IM1EGER(I,J,H,N)
      COMMON /AHCZGK/ AHO, Z, USTART, K
      COMMUN SS
      EXTERNAL FUREK, DFMEK
      SS = 5
      CALL ZERG (FCMEK, DFMEK, OSTART, OMEGA)
      OSTART = UMEGA
      EMEK = DEXP (-OMEGA+Z)/(S+OMEGA) - DEXP(-Z+(AHO+S))/(S+S)
      SETUAN
      END
```

```
C
      DOUBLE PRECISION FUNCTION PREK (5)
      IMPLICIT MEAL * 8 (A-H, K, L, O-Z), INTEGER (I,J, A, A)
      COMMUN /AHOZOK/ MHO, 4, CSTART, K
      FM2K = S-EnG+ (1-(K/(K+S))++K)
      BETUEN
      ENC
C
      DOUBLE PRECISION FUNCTION DEMEK(S)
      IMPLICIT REAL #8 (A-H, K, 1, 0-2), INTEGER (I,J, A, M)
      COMMEN /BHOZUK/ BHO, Z, OSTAR1, K
      DE MEA = 1-800+(A/(K+S)) ++ (K+1)
      AETUAL
       ED D
C
       DOUBLE PARCISION FUNCTION DUFMER (S)
      IMELICIT REAL+8 (A-t, K, L, O-Z), INTEGER (I, J, N, N)
      CUMMUN /MHOZUK/ RHO, Z,OSTART, K
      DDFMEK = kHC*(K+1)*(K/(K+S))**(K+2)/K
      METURN
      EL C
C
       DOUBLE PRECISION FUNCTION FOREK (CHEGA)
      IMPLICIA REAL+8(A-6, K, L, O-2), INTEGER (I,J, M, N)
      COMMUN SS
       FUMEK = FALK (OMEGA) - SS
      METUAN
       END
C
      SUBBOUTINE RZGAM (ITABLE)
       IMPLICIT MEAL+8 (A-b, K, L, O-Z), INTEGER (I,J, M, N)
      COMMON /INOUT/ TVAL (16) , KVAL (7) , SCLUT (10, 16, 3) , RHOIN ,
      125 IN (3)
       CUMMON /LAFLAC/ V (50) , M, N
      COMMCN /MHUZO/ MHO, 2, OSTART
       EXTERNAL PGAM, PGAM, DPGAM
C
       SCALING WITH SIGSUR = RHO
       BHC = MHOIR
       ALFS QA = (1-4HO) + (1-4HO) / HO
       "TECTH = DABS (1-HHC) /AHO
       GSTART = 1
              = ZSIN (ITAELE) / WTFCTE
C
```

```
IF (AHC .LE. 1.6) GO TO 10
      SBAR = RHO - 1
      DELTA = - FGAM (SBAR)
      START = SHAR + DELTA
      CAIL ZERO(FGAM, DFGAM, START, OSTART)
  10 CONTINUE
      DO 20 J=1,16
      T = TVAL (J) /ALFSUR
      IF (T . LE. 2) EZIT = 0.
      IF (2 .GT. 2) CALL LINV (PGAM, N.T. BZIT, V.A)
      EZHT = 2 + (EHO-1.) *T + EZIT
      SCINT (10, J, ITABLE) = BZHT+ WTFCTR
  20
      CONTINUE
      BETULN
      END
C
C
      DOUBLE PRECISION FUNCTION PGAM (5)
      IMPLICIT MEAL+8(A-H, K, L, U-Z), IM1 EGER (I, J, M, N)
      COMMON /RHOZO/ AHC,Z,OSTART
      COMMCH SS
      EXTERNAL FUGAM, DEGAM
      55 = S
      CALL ZERO (FOGAM, DEGAM, OSTART, UMEGA)
      CSTABT = OMEGA
      PGAM = DEXP (-OMEGA+Z) / (S+OMEGA)
      BETUAL
      END
      DOUBLE PRECISION FUNCTION FGAM (S)
      IMPLICIT REAL+8(A-H, K, L, O-Z), INTEGER (I,J, A, N)
      COMMON /RHOZU/ MHC, 2, OSTART
      FGAM = S-AHO+DLOG (1+S)
      AETULE
      ENL
C
C
      DOUBLE PRECISION FUNCTION DEGAM(S)
      IMPLICIT LEAL+8(A-H, K, L, C-Z), INTEGER (I,J, A, L)
      COMMON /RHO 20/ MHG, 2, OSTART
      DFGAM = 1-EHO/(1+5)
      HETUMB
      END
```

```
DOUBLE PARCISION FUNCTION FUGAR (OREGA)
      INTLICIT REAL *8 (A-H,K,L,O-Z), INTEGER (I,J, M, M)
      COMMCA /AHOZO/ AHC, 2, OSTART
      CUMMUN 55
      FOGAN = FGAM (UMEGA) - SS
      BETULN
      ENU
      SUBBOUTIME ZERO (F, CF, XSTART, X4ERC)
      IMPLICIT BEAL *8 (A-H, K, L, O-2), INTEGER (I,J, M, M)
      COMMON /NEWRAP/ TCLENC, MAXITM
      ITEMAT = 0
C
      X = XSTART
  10
      CONTINUE
      XNEN = X - E(X)/DE(X)
      BELACC = (XNEW-X) / XNEW
      X = ANE
      ITERAT = ITERAT + 1
      IF (ITERAT . GE. MAXITN) GO TO 50
      IF (DABS(RELACC) .LE. TOLREC) GO TO 20
      GU 1 C 10
  20
      CONTINUE
      XZEAC = X
      BETULN
     WE ITE (6,60)
  50
     FOLMAT (1H1, " NEWTON-RAPHSON SEARCH IN SUBMOUTINE ZERO",
     1. EXCEEDED HAXIMUM ALLOWABLE NUMBER OF ITERATIONS. 1)
      WHITE (6,62) NAXITH, RELACC, TOLANC
     PO BM A1 (1HO, 'MAXITH =', 14,///, 'RELACC =', D28. 10,
     1///. 'TOLANC =' .D28.16)
      STUP
      EN D
C
      SUBROUTIME AZOUT
      IMPLICIT REAL+8(A-H, K, L, O-Z), INTEGER (I,J,A, h)
      COMMON /INOUT/ TVAL(16), KVAL(7), SCLUT(10,16,3), HHOIN,
     145 IN (3)
      DIMENSION ISCLUT (10,16,3)
      DO YC MCOPY=1,5
      WE ITE (6, 10)
  10 PORMAT (141,///)
      MARGIN = 4 + MCOPY
      DO 12 IMMGN=5, MARGIN
      WR IT & (6,11)
  11 FOABAT (14 )
  12 CONTINUE
```

```
DO 80 17Abl E= 1,3
C
      SCALCA = 1000.
      IF (SCLUT (1, 16, ITABLE) .GE. 10.) SCALOR = 100.
      UG 14 I = 1,10
      DO 14 J = 1,16
      ISCL at (1, J, ITABLE) = SCALOR*SCL at (1, J, ITABLE) + 0.5
      CONTINUE
      IF (SCALOR .EQ. 100.) MAITE (6, 16) RHOIM, &SIN (ITABLE)
     FORMAT (1H-, 10X, 'MHC = ',F5.2,6X, 'SCALED INITIAL SERVER',
     1' LOAL =', F5.2,12%, 'SCALED MEAN WAIT IN HUNDHEDTHS')
      IF (SCALOR .EQ. 1000.) WRITE (6,20) RHOIM, ZSIN (ITABLE)
     FOREAT(1H-, 10X, "HHO = ", F5. 2, 6A, "SCALED INITIAL SERVER",
     1' LOAD =', F5. 2, 11X, 'SCALED MEAN WAIT IN THOUSANDTHS')
C
      WRITE (6,25)
  25 FORMAT (1H ,55%, SCALED EPOCH')
       WRITE (6,30) TVAL
      FOBMAT (1H , 21x, 9 : 5. 2, 7F5. 1)
  30
      white (6,35)
      FORMAT (1H , 21X, 16 (1X, '----'))
  35
      WAITE (6,40) (ISCL WY (1,J, ITABLE), J=1,10)
      FORMAT (1H , 1CX, " # IEN ER
  40
                                   ',16(1X,14))
      WRITE (6,45) (ISCLUT (2,J, ITABLE), J=1,10)
      FORMAT (1H , 10X, 'ME1 QUEUE ', 16 (1X, 14).)
  45
      LU 60 1=3,9
      WEITE (5,50) KVAL (I-2), (ISCLAT(I,J, ITABLE), J=1,16)
      FURNAL (1H , 10X, 'ME1 K=', F5. 2, 16(12, 14))
  20
  00
      CONTINUE
      waite (6,70) (ISCL MT (10, J, ITABLE), J=1,16)
      FUAMAT (16 , 10X, GAMMA INPUT', 16 (12, 14))
      CONTINUE
  60
C
  90
      CONTINUA
      BETUBA
      END
```

```
COMPUTER PROGRAM FOR APPENDIX TABLES, RHO EQUAL TO ONE
       IMPLICIT REAL+8(A-n, A, L, O-Z), INTEGER (I, J, A, N)
       CUMMON /INUUT/ TVAL (10), KVAL (7), WAIT (10,10,3), ZIM (3)
       COMMON /LAPLAC/ V (5G) , M, M
       COMMON /NEWBAP/ TOLENC, MAXITA
       M=U
       N=34
       MAXIT N=20
       TOLENC= 1. D- 20
       DO 60 IPAGE=1,3
       READ (5, 20) (ZIN (LTABLE), ITABLE= 1, 3)
READ (5, 20) (TVAL (J), J=1,6)
READ (5, 20) (TVAL (J), J=7, 12)
READ (5, 20) (TVAL (J), J=13, 16)
  20
       FURNAT (OF10.C)
       DO 50 ITABLE= 1,3
       CALL BZ BNK (ITABLE)
       CALL RZMD1 (ITABLE)
       CALL RZMEK (ITABLE)
       CALL RZGAM (ITABLE)
       CONTINUE
       CALL EZQUT
       CCATINUE
  60
       STOP
       END
C
       ELOCK LATA
       IMPLICIT REAL *8 (A-H, K, L, O-Z), INLEGER (I, J, M, M)
       COMMUN /INOUT/ TV AL (16), KVAL (7), WA IT (10, 10, 3), ZIN (3)
       LATA KVAL/4.,3.,2.,1.5,1.,.5,.2/
       LND
C
C
       SUBBUUTINE LINV (P, N, T, FA, V, A)
       IMPLICIT ABAL #8(A-H,U-Z), INTEGER (1-N)
       COMMON/LAPLAC/EPOCH (10), APPROX (10,6).
       WIMENSIQUE (25), F (25), H (25), V (50)
       IF (M.L... M) GO TO 50
       MH = N/2
       E(1) = 2
       E(2) = 12
       F(1) = 1
       £ (2) = 1
       DO 10 1 = 2, MH
       E(I+1) = (2.DG*(2*I+1)*E(I)) / I
       F(I+1) = F(I) + I
   10
       CONTINUE
       FUE = MH
```

```
H(1) = 2/F(NH)
      LC 2C J = 2, NH
      H(J) = \{(J + + P + B) + E(J)\} / F(H + J + 1)
      CONTINUE
      ISIGN = 2*(NH-2*(NH/2))-1
      DO 40 I = 1, %
      V(I) = 0
      JFIRST = (I+1)/2
      JLAS1 = AINO (I, NII)
      DO 30 J = JFIRST, JLAST
      V(I) = V(I) + H(J)/(F(I-J+1)*F(2*J-I+1))
      CONTINUE
      V(I) = ISIGN*V(I)
      ISIGN = -15IGN
  46
      CONTINUE
      A = A
     FA = 0
  50
      A = .693147 180559945 3094 17232121458D0 / T
      DO 60 I= 1, N
      FA = FA + V(I) *P(I*A)
  60
      CONTINUE
      FA = A+FA
      BETUBN
      END
C
      SUBLCUTINE RZWNE (ITABLE)
      IMPLICIT MEAL *6(A-H, K, L, O-Z), INTEGER (1,J, M, M)
      COMMON /INOUT/ TVAL (16), KVAL (7), WAIT (10,16,3), ZIN (3)
      COMMON /LAPLAC/ V (50) , M, N
      COMMCA /ZO/ Z, OSTANT
      EXTERNAL PWNR
C
              = ZIN (ITABLE)
      DO 20 J=1,16
      T = TVAL (J)
      CALL LINV (PWMR, N, T, TERMS, V, M)
      WAIT (1, J, ITABLE) = 2 + TEAM3
      CUMTINUE
      RETURN
      END
      DOUBLE PARCISION PUNCTION PUNK (S)
      IMPLICIT ALAL+8 (A-H, K, L, O-Z), INTEGER (I, J, M, N)
      CCANON /20/ 2,05TART
      UMEGA = DSQRT (2+5)
      PWME = DEXP (-OMEGA+Z)/(S*OMEGA)
      BETUAL
      END
C
C
```

```
SUBACUTINE BEALT (ITAELE)
      IMPLICIT MEAL *8 (A-H, K, L, O-Z), INTEGER (I,J,M,M)
      CUMMEN /INCUT/ TVAL (10), KVAL (7), WAIT (10, 16, 3), ZIM (3)
      COMMUN /LAPLAC/ V (50) . L. N
      CUMBLE /40/ Z, OSTART
      EXTERNAL PAD1, FAD 1, DEAD1
      USTART = 1
              = ZIN(IZABLE)
      LLZ = DEXF(-4)
C
      DO 40 J=1,16
      T = TVAL (J)
      IF (T . LE. Z) EZIT = 0.
      IF (1 .GT. 2) CALL LINV(PMD1, N, T, HT, V, L)
      IF (T .GT. 2) EZIT = HT + (T-Z) *ELZ
      WAIT (2, J, IT ASLE) = Z + EZIT
  20
      CONTINUE
      BETUAN
      END
C
      DOUBLE PRECISION FUNCTION PMD1 (S)
      IMPLICIT REAL *8 (A-H,K,L,O-Z), INTEGER (I,J, M, N)
      COMMCh /20/ Z, OST ALT
      COMMON SS
      EXTERNAL FOND1, DFMD1
      CALL ZERO (FOND1, DPMD1, OSTART, OMEGA)
      CSTART = ONEGA
      EMD1 = DEXP(-OMEGA+2)/(S+OMEGA) - DEXP(-2+(1+S))/(S+5)
      LETU AN
      END
      DOUBLE PRECISION FUNCTION FAD1 (S)
      IMPLICIT REAL+8(A-H,K,L,O-Z), INTEGER(I,J,A,N)
      COMMCA /ZO/ Z, OSTART
      IMD1 = S - 1. + DEXP (-S)
      BETUAN
      END
      DOUBLE PRECISION FUNCTION DEMD1 (S)
      IMPLICIT REAL+8 (A-b, K, L, O-2), INTEGER (I, M, M, M)
      COMMON /20/ 4,05TART
      DFMD1 = 1.-DEXF(-S)
      KETUAN
      END
      DOUBLE PRECISION FUNCTION FORD 1 (CHEGA)
      IMPLICIT MEAL+8 (A-H, K, L, O-Z), INTEGER (I,J, M, N)
      COMMUN SS
      FOMD1 = FMD1 (OMEGA) - SS
      BETUEN
      EN D
```

```
SUBLOUTINE RAMEK (ITAELE)
      IMPLICIT REAL+8(A-H, K, L, O-Z), INTEGER (I,d, A, N)
      CUMBLA /INOUT/ TVAL(16), KVAL(7), WAIT(10,16,3), ZIN(3)
      COMMON /LAPLAC/ V (50) , A, N
      COMMUN /ZOK/ Z,OSTART, K
      EXTERNAL PAEK, FMEK, DFMEK
      DO 3C I=3,9
              = KVAL (1-2)
      OSTART = 1
             = ZIS (ITABLE)
      EL2 = DEXP(-2*(K+1)/K)
      DO 20 J=1,16
      T = TVAL(J)
      IF (T .LE. 2) EZIT = 0.
      IF (T .GT. Z) CALL LINV(PHEK, M, T, HT, V, H)
      IF ('1 . GT. 2) EZIT = HT + (T-Z) * ELZ
      WAIT (I, J, ITABLE) = 2 + EZIT
  20
      CONTINUE
  30
     CONT INUE
      BETUAN
      END
C
      DOUBLE PARCISION FUNCTION PARK(S)
      IMPLICIT REAL *8(A-H,K,L,O-Z), INTEGER(I,J,M,M)
      COMMON /ZOK/ Z,OSTART, K
      COMMON SS
      EXTERNAL FOMEK, DFMEK
      CAIL ZERC (FCMEK, DFMEK, OSTART, OMEGA)
      OSTART = UMEGA
      PHER = DEXP (-OMEGA+Z)/(S+OMEGA) - DEXP(-2+(S+(K+1)/K))/(5+S)
      RETURN
      END
C
      DOUBLE PRECISION FUNCTION FREK (S)
      IMPLICIT REAL+8(A-H,K,L,O-Z), INTEGER (I,J,M,M)
      COMMON /ZOK/ Z,OSTART, K
      FMEK = S-(K+1)*(1-((K+1)/(K+1+S))**K)/K
      BETUEN
      END
      LOUBLE PRECISION FUNCTION DFREK(S)
      IMPLICIT REAL +8 (A-H, K, L, O-Z)., INTEGER (I, J, M, N)
      COMMUN /ZUK/ Z,GSTAR1,K
      DEHEK = 1- ((K+1)/ (K+1+5)) ** (K+1)
      BETUEN
      END
```

```
LOUBLE PRECISION FUNCTION FOREK (CREGA)
      IMPLICIT ALAL+8(A-H, K, L, 0-Z), IMTEGER (I, J, a, h)
      COMMON SS
      FOMER = FALK (OMEGA) - SS
      BETUAN
      ELD
C
      SUBBOUTINE AZGAM (ITABLE)
      IMPLICIT REAL+8(A-H, K, L, 0-Z), INTEGER (I,J, h, m)
      COMMOL /INOUT/ TVAL (16), KVAL (7), WAIT (10, 16, 3), ZIN (3)
      COHMON /LAPLAC/ V (50), M, N
      COENCY /40/ Z, OSTART
      EXTERNAL PGAM, FGAB, DFGAM
      OSTART = 1
              = ZIN (ITABLE)
C
      DO 20 J=1,16
      1 = 1VAL (J)
      IF (T .LE. 4) E411 = 0.
      IF (T .GT. 2) CALL LINV(PGAB, N. . . EZIT, V. A)
      WAIL (10, J, ITABLE) = 2 + EZIT
  20
      CONTINUE
      RETURN
      END
C
C
      DOUBLE PRECISION FUNCTION PGAM (S)
      IMPLICIT MEAL = 8(A-H, K, L, O-Z), INTEGER (I,J, M, N)
      COMMON /20/ 4,05TART
      COMMUN SS
       EXTERNAL FUGAR, DEGAM
      SS= 5
      CALL ZERO (FOGAM, LEGAM, OSTART, OMEGA)
      CSTART = OMEGA
      PGAM = DEXF (-OMEG A+Z) / (S+OMEGA)
      RETURN
      EL D
C
      DOUBLE PRECISION FUNCTION FGAM (S)
       IMPLICIT REAL+6(A-H, K, L, O-Z), LHTEGER (I, J, M, M)
      COMMON /20/ 2,05TART
      FGAM = S-DLOG (1+5)
      RETURN
       END
C
      DOUBLE FRECISION FUNCTION DEGAM(S)
       IMPLICIT REAL+8(A-H,K,L,O-Z), INTEGER (I,J, M, M)
       COMMON /20/ Z,OSTART
       DFGAL = 5/ (1+5)
       BETUAN
       END
```

```
C
      DOUBLE PARCISION FUNCTION FOGAM (CMEGA)
      IMPLICIT MEAL +8 (A-H, K, L, O-Z) , INTEGER (I, M, B, N)
      COMMON SS
      FUGAL = FGAM (OMEGA) - SS
      BETUEN
      END
C
C
      SUBACUTINE ZERO (F, DF, XSTART, XZERC)
      IMPLICIT HEAL+8(A-H, K, L, 0-2), INTEGER (I, J, M, N)
      COMMON /NEWRAP/ TGLRNC, MAXITM
      ITERAT = 0
C
      X = ASTART
  10
      CONTINUE
      ABEW = \lambda - F(X)/DF(X)
      RELACC = (XNEW-X) / XNEW
      A = ABEN
      ITERAT = ITERAT + 1
      IF (ITERAT .GE. MAXITM) GO TO 50
      IF (DABS(RELACC) .LE. TOLRNC) GO TO 20
      60 TC 10
  20
      CONTINUE
      XZERO = X
       BETUAN
C
  50
      hA ITE (6,60)
  60 FORMAT (1H1, NEWTON-BAPHSON SEARCH IN SUBBOUTINE ZERO",
      1 · EXCEEDED MAXIMUM ALLOWABLE NUMBER OF ITERATIONS. ·)
       WRITE (6,62) MAXITA, RELACC, TOLANC
      FORM AT (1HO, 'MAXITA =', 14,///, 'RELACC =', D28.16,
     1///, 'TOLANC =' , D28.16)
      STOP
      ENL
C
C
      SUBROUTINE BZOUT
      IMPLICIT REAL+8(A-H, K, L, O-Z), INTEGER (I, J, M, N)
      COMMC# /INOUT/ TVAL(16), KVAL(7), WAIT(10,16,3), ZIN(3)
      WIMENSION INALT (10, 16, 3)
C
      DO 90 ACUPY=1,5
      WE IT E (6,10)
      FORMAT (1H1,///)
  10
      MARGIN = 6 + MCOPY
      DO 12 IMAGE =7, MARGIN
      WE ITE (6, 11)
  11
     PORMAT (1H )
      CONTINUE
  14
```

```
DC BU ITABLE= 1,3
      SCALLE = 100C.
      IF (WAID (1, 16, ILABLE) .GE. 10.) SCALOR = 100.
      DO 14 4 = 1,10
      LO 14 J = 1,16
      IMAIT (I,J, ITABLE) = SCALOR* MAIT (I,J, ITABLE) + 0.5
  14
      CONTINUE
      WEITE (6, 25)
       FORM AT (1H-, 16X, PARAMETERS FOR THESE PROCESSES SELECTED',
     1 SO THAT VAR(X(1))=1. AND E(X(1))=0.0, I.E., MHO=1.04)
C
      IF (SCALCR .EQ. 10C.) WRITE(6, 16) ZIN (ITABLE)
      FOREAT (1H , 1CX, MEAN WAIT IN HUM DAEDTHS ...
     143x, 'INITIAL SERVER LOAD = ',F3. 1)
     IF (SCALOR .E. 1000.) WRITE (6,20) ZIN (ITABLE) FORMAT (1H , 10X, MEAN WALF IN THOUSANDTHS.,
     142%, 'INITIAL SEAVER LOAD = ',F3.1)
       WRITE (6,30) TVAL
      FORMAT (1H , 10 X, ' EPOCH: ', 5X, 6F5.2, 8F5.1)
       WRITE (6,35)
  35
      FORMAT (1H , 21X, 16 (1X, '----'))
      WRITE (6,40) (IMAIT(1,J,ITABLE), J=1,16)
  40
      FORMAT (1H , 1CX, " WIEN ER
                                    1,16(1X,14))
C
       WRITE (6,45) (IWAIT (2, J, ITABLE), J=1,16)
      FOREAT (1H , 10X, 'AL1 QUEUE ', 16 (1X, 14))
       DO 60 I=3,9
      WRITE (6,50) KVAL (1-2), (IWAIT (I, J, ITABLE), J=1,16)
      FORMAT (1H , 1CX, 'ME1 K=', F5. 2, 16 (1X, 14))
      CONTINUE
       WRITE (6,70) (IWAI' (10, J, ITABLE), J=1,10)
  70
      FORMAT (1H , 1CX, 'GAMMA IMPUT', 16 (1X, 14))
  60
      CONTINUE
  90
      CONTINUE
       BETURN
      END
```

APPENDIX C

TABLES OF SCALED EXPECTED SERVER LOAD

All tables for a specific value of $\,\rho\,$ are grouped together. The tables are arranged in the order shown in this list. Each page contains results for three values of $\,z^{\star}.$

ρ		z*
2.0		2 2 2 2 2 2 2 2 2
1.5	There are two pages	
1.4	of tables for each of these five values	1.0, 0.8, 0.6
1.3	of ρ .	0.4, 0.2, 0.0
1.2		
1.1		
1.0		200
0.9	There are three	4.0, 3.0, 2.0
0.8	pages of tables	1.0, 0.8, 0.6
0.7	for each of these seven values of ρ .	0.4, 0.2, 0.0
0.6		
0.5		

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	X	-	250	-	_	S	8	0	N	626	10	37	83	7	45	•	45	15
I 3	IEI KE	900	250	416	385	452	580	101	80.0	932	13	1367	1889	2401	3412	4416	6418	8419
	EI K	•	253	-		4	-1	0	-	925	4	35	88	30	9	•	2	0
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WALT	•	1	000	4037	037	035	034	933	031	620	027	MAIT		3.0		200	0 4	3	800	8	5	85	84	4			3.0	1	-	-	-1	0	0 4	1692	0	0	9
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S	-	1		1100	2	20	CI	2	2	2	2	S	•	01.0		30	0	0	0	0	0	0	0	0	v	,	0.10	:	0	7	00	36	00	1002	0	0	0
RHO = 1.20		!	I ENCH	ME1 K= 4.00	E1 K= 3	F1 K= 2.	G1 K= 1.	E1 K= 1.	£1 K= 0.	E1 K= 0.	AMMA IND	RHO = 1.20			TENE	OI OIE	FI KE A	FI KI 3	F1 K= 2.	E1 K= 1.	ME1 K= 1.00	E1 K= 0.	C1 K= 0.	AMMA IND	BHO = 1.20			-	TENER	ממבחב	KI KI 4.	2 1 2	FIKEL	ME1 K= 1.00	E1 K= 0.	11 K= 0.	AMMA IND
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IN THOUSANDTHS	4.0 6.0		200 5200 610	252 6276 637	267 6570 857	565 6568 856	562 6565 856	ARG 5 ARG 0AR	200 0000 0000	250 4550 055	550 6553 855	544 6547 854	36 6541 854	IN THOUSANDTHS		.0 6.0 8.	****	30 6534 853	77 6401 940	200	1 98 1 1 99 99	65 6469 846	62 6465 846	59 6462 846	54 6457 845	AK SARO BAS	30 6442 944	4432 6435 8445	IN THOUSANDTHS		6.0 8.		94 6499 850	38 6442 AF	24 CAN 00	548 1549 83	25 6429 842	21 6424 842	18 6421 842	13 6416 841	04 6407 840	06 6399 80	4389 6391 8392	
WATT	2.0		200	220	166	249	546	444		200	535	529		TIVA		.0	-	808	ARO		000	141	**		437	057	203	2416	WAIT		•	-	472	410	000		000	405	399	395	387	170	2372	
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= 1.20						3.	K= 2.00	- 1 =	-		•	5	Z	= 1.20				KER	THE I			•	# 5.		- "	.0 #	-0 =	MA INPUT	= 1.20				VER.	ういろ	A =	-	•				.0 =	.0 .	TUGNI VA	
OHa				3	"	11	U	W	L	U	11	U	•	OHO				IE	5	i				-	E	5	ū	GAM	RHO				-	0	L	U	11	11	1	W.	ш	u	GAM	

HS		0	0	0	C	0	0	C	0	0	0	1	מ		1	000		0	0	0	0	0	0	0		HN	•	!	0	0	C	0	0	0	C	0	0
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HUND		9	00	0	0	9	9	9	9	9	9	-	3		1	130	7 6	35	P	E	P	30	30	30		251			2	8	2	20	50	20	20	8	50
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WATT		800	0	C	C	0	0	C	C	0	0	*		•	1	100	0	C	0	0	0	0	0	0		IV		-	0	0	0	0	0	0	0	C	0
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CALED	0	00	00	00	00	00	00	000	000	000	00		CALED	•	1	200			000	00	000	00	00	00		CALED			00	00	00	00	00	00	00	00	00
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LOAD	0.70	+	-	-	-	-	-	-	-	-	1		2	-	1	370	- 1	-1		-		-	-	-		LOAD	-		-	-	-	-	-	-	-	-	-
FRVER	9	460	0	0	V	C	V	C	W	0	W		2	69	-	360	20	1	0		W	0		0		EXVER	60		0	W	0	0	0	O	0	0	0
TAL S	0.50	450	5	S	-	1	V		S	450	10		2 7 7	05.0	1	350	0 6	38	O	5	1	80	3	5		IALS	3	-	3	3	5	R.	3	3	3	5	5
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CALED	0.30	430		430	430	430	430	2	430	430	430		1	0.30		330		000	330	330	330	330	330	330		CALED	30	1	230	230	230	230	230	230	230	230	230
v	0.20	42	S	S	N	3	S	10	i	N	N	•		N	1	320	20	10	10	N	S	2	N	N	•	0	N	-	N	O	O	2	N	N	S	N	N
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BHO		-	0		u	W	L		W	MEI	~	2					3 8		1	-	L	w	W	<		D I			-	0	n.	U	B!	u	UI	M	ME

1.10	S	CALED	INI	IAL	SERVER	LOAC	11	CAL ED	0	36	ALED	MEAN	WAIT	r z	THOUSA	DTHS
	0.5	0.10	0.15	0.20	0.30		0	9.0	0.80	c	1.5		•		•	
		!:	1		1;	1	į.	!;	18		i.	13	1	1	i	1
	n	25	0	V	200	29	7	5	20	031	41	50	8	D	01	0
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	7	35		vc	200	25	0	0	00	200	7	100	3	96	DE	0
	שו	25	2	10	2 6	2	200	3	0 0	200	7	700	56	> 6	DW	7
	Y	9		10	2	3	3 6			200	7 6	700	3	3	1) (
-		2	, r	10	2	9	200	200	5	000	7	700	3	0) (7 4
	2	2		10	5	3	2	2	-	0	0	000	3)
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	0.5	2	13	S	30	1402	1508	0	81		2527	3033	4040	5043	7045	904
	S	CALED	LINI	TAL	SERVER	LOAC	0 = 0	0.		SC	ALED	MEAN	WAIT	FZI	HOUSAI	TOTH
	•	•	•	(1		1	7	0							
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	8	8	8	0	-	2	0		3	8	37	8	0	0	6	0
	1	000	2000	00	10	5	-	4	63	8	K	800	3		8	8
	5	006	-	00	10	21	-	4	63	84	35	86	87	187	88	88
	8	006	-	00	10	21	-	4	63	84	35	86	87	187	88	88
	0	006		8	2	2	-	4	63	3	33	86	87	87	88	88
	0	006	-	000	2	2	-	4	63	4	512	8	87	100	8	87
	01			36	2	7	٠,	4	21	000	51	80	8	9	5	8
	nu		- 100			10	٠.	4	200	00	26	90	8	00	10	20
	820	000	950	1001	1104	1208	1313	1418	1627	1835	2348	2857	3865	4868	6871	887
	S	CALED	INI	TAL	SERVER	R LOAD	0 = 0	0	- (S	SALED	MEAN	WAIT	FX	HOUSA	DTH
	•	•	•		•	•	•	ALE	2							
	0000	01.0	61.0	0.50	0000	0.40	0.00	0.00	0.80	1.0	1.3	2.0	9		0	8.0
	R	1	78	-	10	1	8	20	A	9	5	10	-	1	1	
	1	101	16		0 (100	1	10	4	35	0	1-	10	-	14	10
		102	3	. C	10	200	1	10	4	3	0	1	1:		25	1:0
	1	701	3	0	10	FO	4	10	4	2	10	10	7		12	10
	5	200	5	C	-	03	4	24	46	67	0	10	7	-	12	12
	-	200	5	0	-	02	13	24	46	67	10	10	7	-	72	12
	5	200	5	0	-	02	13	2	46		_	0	-		6721	10
	5	100	5	C	-	02	13	24	45	67	18	10	71	-	71	71
	3	100	5	0	~	05	13	2	45	9	18	69	10	-	2	-
	5	100	5	0	-	05	13	2	43	8	28	69	2	-	7	7

à	RHO =	1.10		SCALED	INI	LIAL	SERVER	LOAD		0.	6	SC	ALED	MEAN	WAIT	t z	HOUSA	NOTHS
			0.02	0.0	90.0	0.08	0.10	-	202	0	0.60	6	-	N	3	4.0	0	
	2		3	3	3		1 6	583	647	1 4	2	1 34.3	1550	2500	3614	4610	6623	8624
	-	E	420	100	0			1	M	876	10	32	3	57	59	4595	50	90
2	E1 Kz		42	0	0			-	m	873	10	32	53	5	58	1651	29	29
	~	ë	42	011	0			-	m	873	20	32	53	57	58	4590	30	20
	*	ä	12	011	0			0	m	871	60	E	23	51	58	4588	20	20
	¥ .	=	2	0	0			0	m	871	60	5	53	8	58	4586	20	29
	X	-	42	0	0			0	M	869	60	5	53	28	2	4584	58	58
	× :	•	45	440	0			0	N	867	60	5	25	26	51	4561	58	58
	×	0	75	2	0			0	N	865	60	F	52	20	51	4577	58	58
		2	*	0	0			0	N	963	80	8	27	22	20	101	2	2
•	9			604160			FEDVE					,					- 400	ST. CO.
7	2	:		,		1	CHAE			>.	8	00	U	MEAN			•	-
			6	0.04	0			-	200	16	2				•			
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>	Z		N	265	0	-	U	4	0	-	0	2	46	30	25		83	53
	1		N	253	0	N	-	M	0		6	2	43	18	40	1	30	20
2	*		22	251		-	5	N	0		0	2	43	1	10	4	3	50
	×	3	22	251	0	-	()	N	0		0	2	43	47	48	4	6	64
	*	ö	22	250	0	-	•	Ü	0		0	21	43	11	18	4	10	10
	Z:	-	22	250	0	-	4	N	0		0	5	43	47	48	4	64	0
	X :	-	77	546	0	-	4	N	0			2	45	9	0		01	0
		•	200	24.0	DP	-	4	Ņ-	0		0	25	40	94			9	0
3	AMMA	INPUT	223	246	278	310	341	416	184	743	977	1200	1417	2457	3470	4475	6419	8479
à	. 01	1.10		SCALE	INI O	FIAL !	SERVER	LOAD		0		SC	ALED	MEAN	WAIT	FNI	HOUSA	NOTHS
										J	0							
			0.02	40.0	90.0	0.08	0.10	0.15	0. 20	0.40	0.60	0.8	1.0		3.0	0.4		
			!			1	1	1	1	11	1	18	1	1	1			1
	ENE	2	N	0	N	0	0		0		-	2	2	1	48		0	20
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DEPT. OF OPERATIONS RESEARCH' STANFORD UNIVERSITY, STANFORD, C	ALIF.	10. PROGRAM ELEMENT, PROJECT, TASK AREA & WORK UNIT NUMBERS NR-047-061
OPERATIONS RESEARCH PROGRAM OFFICE OF NAVAL RESEARCH CODE 434 ARLINGTON, VA. 22217		12. REPORT DATE June 1979 13. NUMBER OF PAGES 148
4. MONITORING AGENCY NAME & ADDRESS(II dillorent	from Controlling Office)	18. SECURITY CLASS. (of this report) UNCLASSIFIED
		154. DECLASSIFICATION/DOWNGRADING SCHEDULE

16. DISTRIBUTION STATEMENT (of this Report)

APPROVED FOR PUBLIC RELEASE: DISTRIBUTION UNLIMITED.

17. DISTRIBUTION STATEMENT (of the abstract entered in Block 20, Il different from Report)

18. SUPPLEMENTARY NOTES

This research was supported in part by National Science Foundation Grant ENG 75-14847 Department of Operations Research, Stanford University and issued as Technical Report No. $51\,\text{MH}$

19. KEY WORDS (Continue on reverse side if necessary and identify by block number)

Queueing Theory; M/G/1 Queue; Transient Behavior;

Server Load; Virtual Waiting Time.

20. ABSTACT (Continue on reverse side if necessary and identify by block number)

This research provides numerical results for time-dependent expected server load (mean virtual waiting time) in single-server queues with Poisson arrivals and gamma distributed service times. The results are presented in tabular form to facilitate their use by practitioners involved in the study of operating systems.

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ABSTRACT (continued)

The server load process is expressed as a net input process (having stationary, independent increments) modified by a reflecting barrier at the origin. In queueing systems the net input process is compound Poisson. Two other choices of net input processes, the Wiener process (or Brownian motion) and the gamma process, provide approximations for the queueing process. This research considers groups of server load processes whose parameters are selected so that the first and second moments of their net input processes are matched. An existing Laplace transform expression is employed to obtain transient expected server load at specified epochs.

The Laplace transform is inverted numerically using an existing technique. The inversion technique is first applied to test functions whose exact inverses are known, and the results are also compared with queueing results obtained by other methods. These investigations indicate that the numerical results for expected server load have four to six significant digits when the approximation is based upon thirty four values of the Laplace transform function. The computer programs are coded in FORTRAN IV using extended double precision, and complete documentation is included.

The numerical results for expected server load are tabulated in scaled form. Each of the 93 tables includes results for a specific combination of traffic intensity parameter (twelve values, ranging from .5 through 1.0 up to 2.0) and initial server load (either six or nine values, ranging from 0 to 4 in scaled form). An individual tables has results for the Wiener process, the gamma input process, and eight queues with different service time distributions; each of the ten processes is evaluated at sixteen epochs. Step-by-step procedures for using the tables are included, and several sample problems are presented.

The tabulated results allow a comprehensive study of the error associated with using the Wiener process as an approximation of server load in queues. This study confirms that the Wiener process is always an upper bound and that the approximation is best for queues with a traffic intensity parameter near unity. The scaled results also indicate that the gamma input process and queueing process with deterministic service times provide tight lower and upper bounds, respectively, for expected server load in all queues with Poisson arrivals and gamma distributed service times.